# ACT RESEARCH REPORT





	·				
·					
ALT	FERNATIVES TO	SEX-RESTRICTIV	E VOCATIONAL	INTEREST ASSES	SSMENT

Prepared by the Research and Development Division The American College Testing Program

For additional copies write: ACT Publications P.O. Box 168 Iowa City IA 52243

(Check or money order must accompany request.) Price: \$1.00

© 1979 by The American College Testing Program. All rights reserved.

#### **ABSTRACT**

According to a recent survey, approximately 3,500,000 interest inventories are administered each year by vocational counselors and others in the helping professions. Many of the interest inventories in common use are sex restrictive in that the scores obtained by males typically suggest "man's work" while those obtained by females typically suggest "woman's work." The primary purpose of this report is to summarize recent research on interest assessment procedures as it bears on the issue of sex restrictiveness. Studies documenting sex restrictiveness in widely used interest inventories are cited, and various alternatives to sexrestrictive interest assessment are suggested. Separate consideration is given to basic (homogeneous) interest scales and occupational scales. For basic interest scales, results are summarized for 15 recent validation studies which support the use of interest scores that are not sex restrictive. Such scores can easily be obtained for traditional scales that assess basic interests. Various alternatives to traditional, sex-restrictive occupational scales are reviewed, including the use of cross-sex scales, combined-sex scales, cluster scales, and sex-balanced scales. Normative and correlational data are presented as aids in mitigating sex restrictiveness in existing occupational scales, and the value of placing greater emphasis on basic interest scales is illustrated. For both basic interest scales and occupational scales, primary attention is given to the possibility that, contrary to current practice, interest inventories can consist entirely of items that elicit similar responses from males and females. The scores for such "unisex" inventories would provide males and females with similar vocational suggestions. Reliability and validity data for one such inventory are summarized.

)

	.·

# TABLE OF CONTENTS

Purpose and Scope of the Report	. 1
Basic Vocational Interest Scales	. 2
The Problem of Sex Restrictiveness	
Validation Models  The Comparative Validity of Sex-restrictive and Sex-balanced	. 5
Reporting Procedures	. 6
Must Interest Scores Be Sex Restrictive in Order to Be Valid?	
Sex-balanced (Unisex) Scales	. 11
The Comparative Validity of Sex-balanced (Unisex) Scales and	
Traditional Scales	
Concluding Comments	. 15
Occupational Interest Scales	. 18
The Problem of Sex Restrictiveness	. 18
Alternative Methods of Scale Construction	
Alternative Methods of Interpreting Scores on Existing Scales	
Concluding Comments	. 28
Reference Notes	. 29
References	. 31
Appendix	. 35
Strong-Campbell Interest Inventory Intercorrelation Matrix for Homogeneous and Occupational Scales	

•			
		•	
		•	

#### ALTERNATIVES TO SEX-RESTRICTIVE VOCATIONAL INTEREST ASSESSMENT

#### Dale J. Prediger

#### Richard W. Johnson

Vocational interest inventories have been popular counseling tools for nearly 40 years. According to a recent survey (Tittle & Zytowski, 1978), approximately 3,500,000 interest inventories are administered each year by vocational counselors and others in the helping professions. One of the main applications of interest inventories in vocational counseling is in helping counselees identify vocational options they otherwise might not have considered. Yet, as this report shows, many of the interest inventories in common use are sex restrictive in that the scores obtained by males typically suggest "man's work" while those obtained by females typically suggest "woman's work." Perhaps this should not be surprising, considering that current

interest assessment procedures were developed over 40 years ago. However, the continued use of sex-restrictive interest scores, particularly with females, is problematic. The occupations that such scores suggest to females tend to have lower pay, less responsibility, less status, and fewer opportunities for advancement. Two sets of guidelines on this matter, one by the National Institute of Education (Tittle & Zytowski, 1978) and the other by the Association for Measurement and Evaluation in Guidance (AMEG, 1973), stress that the occupational options suggested to individuals through the use of an interest inventory should not be limited solely on the basis of gender.

### Purpose and Scope of the Report

The primary purpose of this report is to summarize recent research on interest assessment procedures as it bears on the issue of sex restrictiveness, or conversely, sex fairness. Studies documenting sex restrictiveness in current assessment and reporting procedures are cited, and various alternatives to sex-restrictive interest assessment are considered. Recent research, much of which is reported for the first time, is summarized. Major attention is focused on the possibility that, contrary to current practice, interest scales can consist entirely of items that elicit similar responses from males and females. Such scales provide males and females with similar vocational suggestions. Research on the psychometric characteristics of sex-balanced interest scales is summarized.

This report is divided into two sections that correspond to the two main types of interest scales: basic interest scales and occupational scales (Anastasi, 1976). Occupational scales report scores for specific occupations; a separate scale is required for each occupation covered by the interest inventory. The Strong Vocational Interest Blank (SVIB) is a frequently cited example of an interest inventory using occupational scales. Recent revisions of the SVIB, now called the Strong-Campbell Interest Inventory (SCII), have increased the number of occupations for which scales are provided.

Basic interest scales report scores for general types or categories of interests (e.g., social, mechanical, artistic), each of which is relevant to a number of occupations. Occupational relevance is verified by

Dr. Prediger is a research psychologist in the ACT Research and Development Division.

Dr. Johnson is Associate Director of the University Counseling Service at the University of Wisconsin-Madison.

analyzing the scores of persons pursuing various occupations. Perhaps the classic example of an interest inventory with basic interest scales is the Kuder Preference Record—Vocational. But even the SVIB, which is traditionally identified with occupational scales, reports scores for 23 "Basic Interest Scales" and 6 "General Occupational Theme Scales."

Typically, basic interest scales are used to enhance self-knowledge and to suggest career (educational and vocational) options compatible with a person's pattern of interests. They may also reinforce a current choice. The same basic interest scales can serve all three purposes, as can occupational scales. However, it may be difficult to use occupational scales for enhancement of self-knowledge.

This report is based on two symposium papers presented at the 1978 National Convention of the American Psychological Association. The sections on basic interest scales and occupational scales were prepared by Prediger and Johnson, respectively; the report reflects the informal styles of the original convention papers.

#### **Basic Vocational Interest Scales**

In this section, problems related to sex restrictiveness in basic interest scales are discussed, and possible alternatives are considered in the context of recent research. Readers who have followed the literature on sex bias in interest assessment will find that some old problems are addressed once again. Since the "old problems" and "old data" aren't really that old or that widely known, a brief review is provided in conjunction with the results of some recent studies.

## The Problem of Sex Restrictiveness

Current problems of sex restrictiveness in vocational interest assessment appear to rest on the

belief that sex-restrictive interest inventories simply report facts of life. They are Mother Nature's way of reminding vocational counselors that boys and girls are different. However, many counselors may not be aware of what the term "sex restrictive" really means or the degree to which sex restrictiveness exists in widely used interest inventories.

A definition of sex restrictiveness. One way to define sex restrictiveness is through some actual data. According to Gottfredson, Holland, and Gottfredson (1975), for example, the distributions of Self-Directed Search (SDS) high-point codes (highest scores) for "diverse samples of 2,169 high school boys [and] 2,447 high school girls" (p. 139) are as follows:

SDS scale	Girls	Boys	Occupational categories associated with SDS scale
S	67%	20%	Education and social welfare occupations
Α	13	8	Artistic, musical, and literary occupations
С	11	3	Office and clerical occupations
R	1	40	Skilled trades, technical, and some service occupations
1	8	23	Scientific and some technical occupations
Ε	1	6	Managerial and sales occupations

Over 90% of the high school girls receive their highest SDS raw scores for S, A, or C (social, artistic, office, clerical, etc.) occupations. Only about 10% score highest on I, R, or E (scientific, trades/technical, managerial, etc.) occupations. In contrast, 70% of the boys receive their highest scores for these latter occupations.

The above distributions of scores for males and females and the associated career options help define the practical aspects and implications of sex restrictiveness. The more formal definition adopted here reads as follows: "An interest inventory is sex restrictive to the degree that the distribution of career options suggested to males and females is disproportionate. Conversely, an interest inventory is not sex restrictive if each career option covered by the inventory is suggested to similar proportions of males and females" (Prediger & Hanson, 1974, p. 97).

I hasten to emphasize, however, that a sexrestrictive inventory is not necessarily sex biased. The distinction between sex restrictiveness and sex bias is crucial, for, as Holland and others have pointed out (e.g., see Gottfredson, et al., 1975; Holland, 1975b), interest inventories may need to be sex restrictive in order to be valid. This reasoning is the basis for the following definition of sex fairness: "In order for a sex-restrictive inventory to be called sex fair, the publisher must demonstrate that sex restrictiveness is a necessary concomitant of validity as commonly defined" (Prediger & Hanson, 1974, p. 101). Stated another way, if sex restrictiveness' cannot be justified on the basis of validity evidence, then it is synonymous with sex bias. Thus, the definition follows principles underlying Equal Employment Opportunity Commission Guidelines (1970). The burden of proof, however, is on the test publisher, not the test user.

Because alternatives would be of little use if sex restrictiveness is a necessary concomitant of validity, research bearing on this issue is the major focus of this paper. However, to further illustrate the extent of the problem posed by sex restrictiveness, it will be useful to look at the degree to which sex restrictiveness is present in various interest inventories.

Incidence of sex restrictiveness. Although the male-female distributions provided by Holland are seldom available for interest inventories, one can

frequently find score means and standard deviations for males and for females. Given these data and the procedure developed by Tilton (1937), it is a simple matter to determine the degree to which the scores of males and females overlap. Table 1 provides illustrative data for scales assessing common dimensions of interests as described by Holland (1973).

Dunnette (1966) has suggested that two distributions differ in meaningful ways if overlap is less than 75 percent. Strong (1955) proposed that "two groups that overlap less than 80 percent are different enough to be considered practically different." (p. 22). Whether or not these criteria are applied, it is readily evident from Table 1 that malefemale score differences on certain scales are substantial. Furthermore, the male-female score differences illustrated in Table 1 are not limited to scales assessing Holland types. For example, male-female score overlap ranges from 46% to 99% (median of 80%) for the 23 SCII Basic Interest Scales, according to data for the General Reference Sample (Campbell, 1977, p. 38). Male-female overlap for 5 scales falls below 75%. Readers are reminded that these data are not necessarily indicative of sex bias.

It is clear from Table 1 and the previous SDS distributions that substantial discrepancies in the scores of males and females are common to traditional interest inventories reporting raw scores or scores based on combined-sex norms. Raw scores, of course, reflect any differences in the responses of males and females to specific items. Since combined-sex norms merely anchor raw score scales to parameters based on the total group of males and females, any sex differences on a raw score scale will be reflected in a scale based on combined-sex norms. Thus, as shown by Cole and Hanson (1975), standard scores based on combined-sex norms produce sex differences similar to those observed for raw scores.

When score reports are based on same-sex norms, however, males and females receive highly similar (sex-balanced) interest profiles and distributions of career suggestions (Cole and Hanson, 1975; Gottfredson, et al., 1975; Prediger and Hanson, 1974). This occurs because of the very nature of the norming procedure, regardless of sex differences in raw score distributions. The Kuder Preference Record—Vocational is a classic example of an inventory using same-sex norms.

TABLE 1

Overlap of Scores for Males and Females on Various Interest Scales Assessing Holland Types

	Scales based on traditional items								
Scale	SD\$ <sup>a</sup>	VPI <sup>b</sup>	SCII <sup>c</sup>	CAId	CDM <sup>e</sup>	ACT-IV <sup>f</sup>	Brand X <sup>9</sup>		
Investigative	77%	85%	88%	90%	91%	84%	93%		
Artistic	78	<b>7</b> 7	75	77	<b>7</b> 7	76	87		
Social	50	62	90	82	56	60	85		
Enterprising	87	90	85	97	86	98	99		
Conventional	75	94	99	74	98	95	97		
Realistic	32	62	65	63	54	57	89		

Note. Percent overlap is based on Dunnette's (1966) table for Tilton's (1937) measure of overlap.

<sup>&</sup>lt;sup>a</sup>Data are based on Self-Directed Search (SDS) summary scores for 2,152 male and 2,431 female high school students (Gottfredson & Holland, 1975a).

<sup>&</sup>lt;sup>b</sup>Data are based on Vocational Preference Inventory (VPI) raw scores for 6,290 male and 6,143 female entering college students (Holland, 1975a, p. 29).

<sup>&</sup>lt;sup>C</sup>Data are based on Strong-Campbell Interest Inventory (SCII) Theme Scales standard scores for 300 males and 300 females in the men- and women-in-general samples (Campbell, 1977, p. 33).

<sup>&</sup>lt;sup>d</sup>Data are based on Career Assessment Inventory (CAI) Theme Scale standard scores for a "composite reference sample" of 750 males and 750 females (Johansson, 1976, p. 23). This sample was used to select a subset of CAI items that minimized theme scale sex differences (Johansson, 1976, p. 20).

<sup>&</sup>lt;sup>e</sup>Data are based on Harrington/O'Shea System for Career Decision Making (CDM) raw scores for 435 male and 380 female high school and college students (Harrington & O'Shea, 1976, p. 9).

<sup>&</sup>lt;sup>1</sup>Data are based on ACT Interest Inventory (ACT-IV) raw scores for the 1,233 males and 1,738 females in the ACT-IV national norm group for college-bound persons (Hanson, 1974, p. 14). These data are for purposes of comparison only. Standard scores based on same-sex norms are used in ACT-IV score reports (Hanson, 1974).

<sup>&</sup>lt;sup>9</sup>Data for 1,247 males and 1,693 females are for a new unisex interest inventory based on sex-balanced items. Brand X data are provided for perspective only.

Overview of alternatives. One alternative to sexrestrictive interest reports, then, involves the use of same-sex norms. Another alternative involves the elimination of sex differences at the item level, as suggested by Harmon (1975). Although the fact is not well known, interest inventory authors have written substantial numbers of sex-balanced items in the past. Responses to about half of the items in current inventories are approximately sex-balanced (Campbell, 1977; Harmon, 1975; Johansson, 1976). The implication is that it may be possible to develop interest inventory scales consisting entirely of sexbalanced items. In that case, both raw scores and standard scores based on combined-sex norms would be sex balanced.

As noted previously, the crucial question with respect to both alternatives to sex-restrictive reports (i.e., use of same-sex norms or sex-balanced items) is whether interest scores must be sex restrictive in order to be valid; or conversely, whether sex-balanced reports have less validity. This question needs to be addressed in order to evaluate alternatives to current practice. First, however, careful attention must be given to the procedures used to determine "validity."

#### Validation Models

Because I believe it is crucial to distinguish among various validation models if we are to make progress in eliminating sex bias from interest assessment, and because I don't know how to explain it any better, the passages that follow were taken more or less intact from a recent article in Applied Psychological Measurement (Prediger, 1977).

As Kuder (1970) noted, "the problem of establishing validity for counseling purposes becomes one of classification; [hence] one of the fundamental questions in judging a vocational interest inventory is how well it differentiates among the specific occupational groups for which it is scored" (p. 209). Strong (1943), although primarily concerned with the differentiation of occupational groups from men or women in general, also recognized the need to differentiate among the occupational groups themselves. Though other, perhaps better, approaches to validation are possible (e.g., determination of correlations with satisfaction or success), interest inventory construction and validation studies have typically focused on criterion group differentiation/classification. . . .

It is well known that the validity of a measuring instrument depends on the purposes for which it is used. Hence, before studying validity, one must ask, "Validity for what?" Interest inventories are commonly used to suggest possible occupational options to counselees. Yet, the validity of inventories is

often reported in terms of their ability to predict future occupational preferences or occupational entry [e.g., see Gottfredson and Holland, 1975b]. As Berdie (1970) has noted, few counselors are interested in predicting whether a counselee will enter (or prefer) occupation A or occupation B. Hence, validity data for this use of interest inventories may provide a distorted view of validity for more common uses. Some of the reasons are discussed below.

#### The "Will-Prefer-or-Enter" Criterion

When predicting the occupations persons will prefer or enter, the nature of employment distributions as well as the nature of occupational preferences must be taken into account. Stated another way, if an interest inventory is to provide accurate predictions of eventual employment, the predictions must accurately reflect the size of each occupational criterion group. To the degree that group membership predictions depart from group base rates, the inventory's predictive accuracy will be lowered.

Interest inventories predicting that persons will enter or prefer occupations in the same proportions as in the past should do well under this approach to validation. For a multitude of reasons (e.g., social expectations, local labor market needs, the contingencies of life), people will continue to state preferences for and enter traditional occupations. Unfortunately, the number of persons in various occupations and occupational preference groups differs widely from group to group (Gottfredson, Holland, & Gottfredson, 1975; Prediger, Roth, & Noeth, 1974). Since the predictions used in validation studies are based on the same scores counselees receive, the occupational options suggested to counseless will reflect the same differences in base rates as the predictions. Under this approach to validation, a "valid" interest inventory in the 1850s would have suggested farming to nearly everyone. The employment status quo will be reflected and reinforced by interest inventories validated in this

# The "Should-Consider" Criterion

The alternative approach to the use of occupational preference and membership as criteria in validating interest inventories assumes that the purpose of interest inventories is to identify career options for counselees to consider rather than to predict the occupations counselees will prefer or enter. To achieve the former objective, an interest inventory must assess the correspondence between a counselee's interests and the interests associated with various occupational groups-regardless of the group base rates. If a counselee's interests are similar to the interests of persons in a given occupation, one would suggest that the counselee consider the occupation, even though relatively few persons are employed in the occupation. The emphasis is on "should consider," not "will enter or prefer." The underlying assumption is that employment data may play an important role in career counseling, but they should not influence interest score reports.

Studies following this approach to interest inventory validation will treat occupational criterion groups (or preference groups) as if they were of equal size. One would expect an interest inventory to suggest engineering to a large proportion of criterion

group members in engineering, nursing to a large proportion of nurses, retail sales to retail sales clerks, horseshoeing to horseshoers, and so on for each of the criterion groups available. The fact that there are relatively few horseshoers in comparison to retail sales clerks is irrelevant. The question asked in this validation analysis is "What proportion of the members of each criterion group would have been asked to look into their occupation by this interest inventory?" Stated differently, the question is "What is the hit rate for each criterion group?" A high hit rate depends on an inventory's ability to differentiate the criterion groups and, thus, minimize the misassignment of members of each of the groups.

In this approach to validation, an interest inventory does not have to suggest retail sales to more counselees than horse-shoeing because there are more retail sales clerks than horse-shoers. "Predictions" are simply based on whichever criterion group a person resembles most. There is no premium placed on providing interest-score distributions that parallel preference or employment distributions. This proposed validation strategy recognizes that, for a number of very practical reasons, many persons may not enter the occupations suggested ("predicted") by an interest inventory.

#### How Choice of Criterion Affects Career Guidance

Perhaps the following example will bring differences between the two approaches to validation into sharper focus. Suppose that in a society built on the caste system, an interest inventory was designed to have high validity in predicting occupational entry. The inventory would suggest few, if any, occupations that were not traditional for a person's caste. To do otherwise would lower its validity. On the other hand, suppose the inventory was designed to identify occupational options compatible with a person's interests-regardless of the proscriptions of the society. Such an inventory may suggest many occupations not traditional for members of the caste. As a result, it would be a poor predictor of occupational entry. Yet, it may do an excellent job of determining occupational compatibility. Even in a time of social change, the score reports might be unsettling, but they could provide beneficial information, both to the individual and to the society (pp. 275-277).

Although useful in some types of research, interest inventories designed to predict which persons will prefer or enter a given occupation present special problems for vocational counseling. In effect, the rationale underlying such inventories says "Cindy may have interests like an engineer and Mike may have interests like a nurse. But few females or males are likely to enter those nontraditional occupations. So let the predictions (score profiles) take into account the relative numbers of males and females who have entered various occupations in the past. In the long run, a higher hit rate will be obtained and the inventory will appear to be more valid." When used in vocational counseling, inventories based on this rationale will reinforce society's occupational sex-role stereotypes and thus further institutionalize the channeling. At first glance, such inventories may appear to have higher validity than inventories designed to report occupational options compatible with a person's interests. But this may be true only if one's purpose in assessing interests is to predict the occupations counselees will enter (or prefer).

Prediger and Cole (1975) provide an extended discussion of this topic as it applies to career counseling and nontraditional occupations for males and females. Prediger (1977) discusses specific implications for validation procedures.

The Comparative Validity of Sex-restrictive and Sex-balanced Reporting Procedures

Now, I would like to return to the key question posed earlier—"Must vocational interest reports be sex restrictive in order to be valid?" Table 2 summarizes the results of 10 studies comparing the criterion-related and construct validity of sexrestrictive and sex-balanced reporting procedures. In each of the studies, sex-balanced reporting procedures were based on same-sex norms. The results cited for sex-restrictive reporting procedures were obtained with raw scores. (As previously noted, combined-sex norms reflect essentially the same male-female differences as raw scores.) All studies used measures of Holland types; and for a given study, both sex-restrictive and sex-balanced reports were obtained from the same interest inventory. Thus, any differences in validity reflect differences in the presence or absence of sex restrictiveness in the reporting procedures.

Six of the studies have appeared in the professional literature and citations appear in the left hand column of Table 2. The other four studies, which were completed during the past year, are described below.

Study 1: procedures. Study 1 involved 11,395 college seniors (5,846 males and 5,549 females) enrolled in 16 major universities located primarily in the midwestern, southern, and southwestern regions of the country. Fifteen states were represented. A high percentage of the 1974-75 incoming freshmen at each of the institutions had completed the ACT Assessment Program (AAP) battery as college-bound students in 1973-74, the first year the ACT Interest Inventory (ACT-IV) was included in the AAP. A roster of 1977-78 seniors was

TABLE 2
Summary of Validity Data for Sex-restrictive and Sex-balanced
Score Reports of Holland Types

Study	Type of validity	Time interval	Sample; No. of males (M) & females (F)	Criterion; No. of criterion groups	Relative performance of sex-balanced reports (SBR) & sex-restrictive reports (SRR)
Prediger & Hanson (1976)	Construct	Concurrent & longitudinal (5 years)	Young adults & adults in 3 samples; M=20,000, F=19,000	Occ. status (2 samples) & preference; M=104, F=104	SBR more in agreement with congruency principle and occupational typology in Holland's theory of careers
Prediger (1976)	Construct	Concurrent	High school & college students & adults in 7 samples; M=18,000, F=20,000	NA	SBR more in agreement with consistency principle in Holland's theory of careers
Prediger & Hanson (1977)	Criterion- related	Concurrent	College seniors; M=5,500, F=5,000	College major; M=5, F=5 (by Holland type)	SBR and SRR hit rate similar a for males; SBR better for females
Hanson, Noeth, & Prediger (1977)	n, Criterion- Longitudinal Young adults; & related (5 years) M=648, F=425 er n. Criterion- Longitudinal College sophomores			Occ. status; M=6, F=5 (by Holland type)	SBR hit rates better for males and females
Hanson, Noeth, & Prediger (1977)	Criterion- Longitudinal College sophomores; related (2 years) M=549, F=894		College major; M=5, F=5 (by Holland type)	SBR and SRR hit rates similar for males; SBR better for females	
Prediger (1977)	Criterion- related	Longitudinal (1-3 years)	College students; F=989	Occ. preference; F=5 (by Holland type)	SBR and SRR hit rates similar for females; SBR data not available for males
1 <sup>b</sup>	Criterion- related	Longitudinal (4 years)	College seniors; M=5,846, F=5,549	College major; M=6, F=6 (by Holland type)	SBR and SRR hit rates similar for males and for females; differences favored SBR
2 <sup>b</sup>	Criterion- related	Concurrent	College seniors; M=929, F=1,033	College major; M=6, F=6 (by Holland type)	SBR and SRR hit rates similar for males and for females; differences favored SBR
3 <sup>b</sup>	Criterion- related	Concurrent	College-bound students: M=737, F=852	Occ. preference; M=6, F=6 (by Holland type)	SBR and SRR hit rates similar for males and for females
4 <sup>b</sup>	Criterion- related	Concurrent	Adults; M=289, F=428	Occ. status; M=14, F=20 (by Holland type)	SBR and SRR match between Holland type of criterion group and highest interest scale mean for group members was similar for males; for females, differences favored SBR

Note. Sex-balanced reports (SBR) based on same-sex norms are compared with sex-restrictive reports (SRR) based on raw scores for the *same interest inventory*. All **s**tudies involved traditional interest items assessing Holland's six types.

<sup>&</sup>lt;sup>a</sup>When SBR and SRR criterion group hit rates differed by less than 5% (e.g., 46% vs. 42%), they were considered to be similar.

<sup>&</sup>lt;sup>b</sup>See descriptions of Studies 1, 2, 3, and 4 in this paper.

obtained from each of the institutions and matched against a roster including the ACT-IV scores of 1974-75 enrollees. Since the college majors of the seniors were known, it was possible to determine how seniors majoring in various areas had scored on the ACT-IV four years earlier.

Because all students had achieved senior status in college less than four years after entry, very little time could have been lost due to dissatisfaction with major or unsatisfactory academic performance. Thus, the study design included an indirect criterion group screen for success and satisfaction. The percentage of ACT-tested enrollees at each of the institutions ranged from 64% to 100% (median of 88%). Hence, a high proportion of all seniors meeting the four-year screen had taken the ACT-IV.

College majors were allocated to Holland types on the basis of the classification system and associated alphabetical index provided by Holland (1972). Interest profiles were allocated to Holland types on the basis of the student's highest score (high-point code) for each of the two types of reporting procedure. All score ties were broken randomly. Thus, each college senior in the sample was allocated to a Holland type by two methods: (a) on the basis of academic major, and (b) on the basis of ACT-IV scores. The former method established the student's criterion group membership. Correspondence between a student's criterion group and high-point code was then determined separately for the sex-restrictive and sexbalanced reporting procedures. "Hit rates" were tallied separately for males and females in each criterion group.

Study 2: procedures. Study 2 involved a subset of 2,096 college seniors in Study 1 plus an additional sample of 903 seniors who had not taken the ACT-IV four years earlier. (Prior ACT-IV data were not needed in Study 2 because the design was cross-sectional rather than longitudinal.) Altogether, there were 2,999 students in the study. The sampling plan involved the random selection of approximately equal numbers of males and females majoring in each of ten fields (e.g., engineering, art, physical sciences). The fields were selected to span Holland's six types. The additional sample of students not in Study 1 was needed to assure sufficient numbers in each type to support separate hit rate analyses for males and females.

The 2,999 students in Study 2 were randomly allocated to two subsamples. The first was asked by mail to complete the ACT-IV and the recently developed Unisex Edition of the ACT Interest Inventory (UNIACT), with items sequenced in that order. The second sample was asked to complete UNIACT and then the ACT-IV. Thus, the administration design was counterbalanced. Both local and home addresses were obtained from the universities. Home addresses were used when local addresses proved to be invalid.

After two follow-up mailings, the last about ten weeks after the initial mailing, usable replies were received from 1,988 of the 2,905 students for whom there was no evidence of bad addresses (68% response rate). The response rate for the total sample was 66%. All students received a four-page report of vocational interests, including reference material, in return for their participation in the study.

A number of students in various engineering and agricultural specialities who had been assigned to Holland's realistic type when the sample was initially selected were subsequently shifted to the investigative type upon final assignment. These and a few other changes, including the elimination of 26 cases with very general majors, were made, to achieve close correspondence with Holland's classification system. As in Study 1, students were also assigned to Holland types on the basis of highpoint interest codes. Correspondence between college major type (i.e., criterion group) and interest type was then determined for the sex-restrictive and sex-balanced scores.

Study 3: procedures. Study 3 involved 2,013 of the approximately 127,000 college-bound students who registered for the October 1977 AAP national test date. Only those students who were high school seniors planning to enroll in college the following fall and who were "fairly sure" or "very sure" of their first occupational choice were eligible for the study. The Study 3 sample was selected from this pool on the basis of the general correspondence of expressed vocational choice, as recorded on the AAP registration sheet, to the six Holland types. The sampling plan involved the random selection of approximately equal numbers of males and females, of each Holland type (i.e., criterion group), with some oversampling in the artistic and realistic categories.

As in Study 2, students in Study 3 were asked by mail to complete the ACT-IV and UNIACT. A counterbalanced design was used. After two follow-up mailings, the last about three weeks after the initial mailing, usable replies were received from 1,589 of the initial sample of 2,013, a 79% response rate. All students received a one-page computer-printed report of vocational interests, supplemented by reference materials.

As in Study 2, there were some changes made in the initial criterion group allocations in order to achieve close correspondence with Holland's classification system. Criterion group hit rates were determined using procedures described for Studies 1 and 2.

Studies 1, 2, and 3: summary of results. The unweighted average hit rates shown in Table 3 provide a general index of the criterion-related validity of the sex-restrictive and sex-balanced reporting procedures used in Studies 1, 2 and 3. This hit rate index treats the criterion groups as being of equal importance in vocational counseling and is appropriate to the "should consider" validation model (Prediger, 1977) described earlier in this paper.

In all three studies, overall hit rates for the sexbalanced and sex-restrictive score reports were similar, with small differences generally favoring sex-balanced reports. Overall hit rates for males and females differed somewhat; however, the male and female criterion groups were not comparable because of differences in the mix of majors (Studies 1 and 2) and occupational preferences (Study 3). Across the three studies, sex-balanced reports produced higher hit rates than sex-restrictive reports for 13 of the 18 female groups and 9 of the 18 male groups.

Study 4: procedures and results. Walsh and his students at Ohio State have conducted a series of six studies in which the SDS was administered to predefined criterion groups (Bingham & Walsh, 1978; Fishburne & Walsh, 1976; Horton & Walsh, 1976; Matthews & Walsh, 1978; O'Brien & Walsh, 1976; Spokane & Walsh, 1978). Across the six studies, there were 717 persons in 34 occupational groups—20 groups composed of females (N=428) and 14 groups composed of males (N=289). With one exception, each study followed the same model. Adults in six occupations representing

Holland's (1973) six types completed the SDS. Raw score means on each of the six SDS summary scales were then calculated for members of each occupation. The one exception (Spokane and Walsh, 1978) involved four groups of adults in two occupations representing Holland's enterprising type.

In their six reports, Walsh et al. place major emphasis on a comparison of the means obtained by all six occupational groups on a given SDS summary scale. However, results are also reported for an "intraoccupational analysis" which follows Holland's (1973) proposal for determining, empirically, high-point codes (Holland types) for occupations. Raw score means for all six SDS summary scales were ranked for a given occupation, and the scale with the highest mean was determined. This scale, which constitutes the occupation's empirically derived high-point code, is expected to correspond to the Holland type assigned to the occupation by Holland (1973).

In order to determine the effectiveness of sexbalanced interest reports using the same correspondence criterion of validity, I sought same-sex norms that could be used to convert the SDS summary scale raw scores to standard scores. Of the normative data provided by Gottfredson and Holland (1975a) for adults (N=140), college students (N=3,355), and high school students (N=4,675), the high school data appear to be the most comprehensive. Gottfredson and Holland note that the norms "usually represented large and diverse samples from any sources, but the high school and college samples are believed to be reasonably representative of groups commonly taking the SDS" (1975a, p. 2). Raw score means for the SDS were transformed to standard score means using the high school norms.

Appropriate high-point codes (Holland types) were obtained for 9 of the 20 female occupational groups (45%) when raw scores were used, and for 13 of 20 groups (65%) when standard scores were used. Of the 11 occupational groups with inappropriate raw score codes, 9 scored highest on the SDS Social Scale. Results for males were less discrepant. Raw scores and standard scores produced 7 and 8 appropriate codes, respectively, for the 14 occupational groups. Across all 34 groups, raw scores produced 16 appropriate codes (47%) and standard scores produced 21 appropriate codes (62%).

TABLE 3

Criterion Group Hit Rates for Sex-restrictive and Sex-balanced Score Reports of Holland Types

Criterion group						
Criterion group by Holland type	Sample size			strictive orts <sup>a</sup>	Sex-balanced reports <sup>b</sup>	
	М	F	M	F	М	F
			Study 1			
Investigative	2,008	999	66	42	51	54
Artistic	778	1,353	28	36	47	42
Social	621	2,343	46	76	2 <b>6</b>	22
Enterprising	1,147	414	25	11	32	29
Conventional	457	297	28	32	48	41
Realistic	835	143	31	2	38	27
Unweighted average hit rate			37	33	40	36
			Study 2			
Investigative	323	348	59	49	46	50
Artistic	148	188	62	60	79	57
Social	151	182	41	56	30	24
Enterprising	121	121	57	36	62	43
Conventional	105	118	31	51	47	55
Realistic	81	76	37	9	33	47
Unweighted						
average hit rate			48	44	50	46
			Study 3			
Investigative	187	181	50	28	40	32
Artistic	142	187	40	46	55	41
Social	76	132	54	62	41	26
Enterprising	124	145	44	22	37	34
Conventional	101	132	50	42	69	62
Realistic	107	75	42	9	36	35
Unweighted average hit rate			47	35	46	38

Note. All studies involved traditional interest items assessing Holland's (1973) six types. The same interest inventory was used in each comparison of reporting procedures.

<sup>&</sup>lt;sup>a</sup>Reports are based on raw scores.

<sup>&</sup>lt;sup>b</sup>Reports are based on standard scores derived from same-sex norms.

These results imply that sex-balanced SDS standard scores are more accurate than the sex-restrictive raw scores in describing the Holland types characterizing various occupational groups and, hence, individuals in those groups.

Must Interest Scores Be Sex Restrictive in Order to Be Valid?

The previous section summarizes results for ten studies comparing the construct and criterionrelated validity of sex-restrictive and sex-balanced procedures for reporting interest scores. Included were concurrent and longitudinal studies involving a variety of criterion groups and instruments (e.g., the SDS, VPI, SVIB, and ACT-IV). Results from each of the studies indicate that the validity of interest inventories is not lowered through the use of sex-balanced score reports. In several instances, it is increased. A recent study by Lamb (1975) also indicates that sex-balanced score reports are appropriate for use with males and females in various minority groups, and several studies show that persons in a wide range of criterion groups obtain sensible score profiles when sex-balanced reports are used. Hanson, Prediger, and Schussel (1977) use high-point codes based on same-sex norms to summarize longitudinal and crosssectional data on the vocational interests of 103 educational criterion groups (N=18,435), occupational criterion groups (N=1,073) and 39 vocational choice groups (N=7,148). The threeletter codes for males and females are generally congruent with expectations based on Holland's theory of careers.

In six of the ten studies summarized in Table 2, sexrestrictive reports based on combined-sex norms were included in the comparisons. (Results were not summarized to avoid complicating the table.) The validity of these reports was sometimes higher than the validity of the sex-restrictive raw scores summarized in Table 2, but it in no case exceeded that of the sex-balanced reports. One other study relevant to this issue (Gottfredson & Holland, 1975b) is sometimes cited as showing that sexbalanced reports are "less valid." In that study, sexrestrictive reports did produce more accurate predictions of future vocational preference for college women. As already noted, however, this approach to validation (i.e., prediction of future preference) is not applicable to counseling uses of interest inventories. In any case, the predictions failed to improve upon the base rates or predictions based on current preference—alternative procedures for predicting future preference that ignore interest scores.

Considered as a whole, the validity data summarized in Table 2 indicate that vocational interest inventories do not have to be sex restrictive in order to be valid. Thus, sex-restrictive interest score reports may well be sex biased. Interest inventory publishers and other researchers might consider conducting additional studies to compare the validity of sex-restrictive and sex-balanced reporting procedures.

# Sex-balanced (Unisex) Scales

Rationale. The data thus far indicate that sex-balanced interest reports based on same-sex norms provide a viable alternative to sex-restrictive reports. As noted previously, another alternative is to eliminate sex differences at the item level and, thus, produce "unisex scales" (Rayman, 1976) based on sex-balanced items. Since males and females will obtain similar scores on these scales, combined-sex norms could be used without being sex restrictive. Same-sex norms which, according to some, "treat males and females differently" would no longer be needed.

Studies by Boyd (1976), Gottfredson (1976), and Holland and Gottfredson (1976) show that simply desexing existing items has little effect on scale scores. However, no attempt was made in those studies to write and pretest new items endorsed in equal proportions by males and females. Rayman (1976), working with Hanson and Cole at ACT, recently demonstrated the viability of this latter approach to interest scale construction. Subsequently, Hanson and Rayman (1976) showed that Rayman's "unisex scales" had criterion-related validity equivalent to that of sex-restrictive scales administered to the same sample.

Encouraged by these results and the related work of Lunneborg (1977), staff members at ACT conducted a series of studies leading to the development of the Unisex Edition of the ACT-IV (UNIACT). Starting with a substantial pool of sexbalanced items already used in various forms of the ACT-IV, we attempted to write additional items that captured the essence of a work-related activity preference white minimizing sex-role connotations. As noted by Prediger and Hanson (1978),

"this approach to interest scale construction recognizes that sex differences in the responses to many interest items may reflect the differential effects of sex-role socialization on males and females without necessarily reflecting differences in basic interests. Thus, groups of males and females may respond quite differently to interest inventory items with sex-role connotations. ..even though the groups may have similar patterns of interests."

Consider, for example, the following interest items which are typical of those included on some "mechanical" or "realistic" interest scales. "Would you like to be—

- a car mechanic?
- a bulldozer operator?
- a train engineer?
- a power shovel controller?"

Should one take sex differences on a scale containing such items at face value—as indicators of fundamental differences in the mechanical (realistic) interests of males and females? Certainly, the items appear to fit the "mechanical" category; they are gender neutral; and they may correlate with other items in the appropriate manner (similar items have). But do they register the interests of males and females in equal measure? Would the scores of males and females differ by the same amount if items free from sex-role connotations were used? The work of Rayman (1976) clearly suggested that the answer is "no." We proceeded on that basis.

The construction of UNIACT primarily involved the development and repeated tryout of potentially sexbalanced items. As described by Hanson, et al. (1977), more than 200 items were used in a series of studies involving six samples (N=10,388) of 9th graders, 11th graders, college-bound students, college sophomores, and adults. Initial data on the psychometric characteristics of UNIACT were reported by Hanson, et al. (1977) and summarized by Prediger and Hanson (1978). Beginning in the fall of 1977, UNIACT replaced the ACT-IV as a core component of the ACT Assessment Program (the "ACT"). A 60-item version of UNIACT is also used in the Vocational Interest, Experience, and Skill Assessment (ACT, 1976). In both assessment programs, UNIACT score reports are based on combined-sex norms.

Psychometric characteristics of sex-balanced scales. Internal consistency reliabilities for the six 15-item UNIACT scales range from .85 to .92, with a

median of .87 for a holdout sample of 914 males and 937 females (Hanson, et al., 1977). Test-retest stability coefficients for about a six-week interval range from .79 to .87 (median .82) for the students in Study 3. Male-female score overlap is shown in Table 4 for four samples. For the UNIACT norm group, the range is 85% to 99% with a median of 91%. Across the four samples, overlap for the Social Service Scale was generally the lowest, yet it ranged from .84 to .85, well above Dunnette's (1966) 75% criterion cited earlier.

As previously noted, UNIACT and earlier editions of the ACT-IV were constructed to assess Holland types. (ACT-IV profiles generally in accord with Holland's theory of careers are provided by Hanson, et al., 1977, for a wide variety of criterion groups.) As reported by Hanson, et al. (1977), the correlations between parallel UNIACT and ACT-IV scales range from .76 to .86, with a median of .80 for a grade 11 sample. In Study 2, correlations for males ranged from .71 to .92 (median of .88). For females the correlations ranged from .75 to .91 (median of .87). Correlations based on Study 3 data for males ranged from .84 to .95 (median of .91). The correlations for females ranged from .80 to .94 (median of .91). As indirect indicators of UNIACT construct validity, these correlations compare quite favorably to the median VPI-SDS correlations of .55 for males and .43 for females reported by Holland (1972), especially since the VPI is a major component of the SDS and both were constructed to assess Holland types (Holland, 1973).

Additional data bearing on the construct validity of UNIACT scales are provided by their factor structure. As shown by the factor loadings in Figure 1, the hexagonal configuration basic to Holland's theory is present for both males and females. (Hanson, et al, 1977, provide further data on this point.)

Both the hexagonal configuration of factor loadings and the correlations between parallel UNIACT and ACT-IV scales suggest that the sex-balanced scales have good construct validity as measures of Holland types.

The Comparative Validity of Sex-balanced (Unisex) Scales and Traditional Scales

As previously noted, sex-balanced vocational interest scales provide an alternative to traditional sex-restrictive scales. However, some have claimed that

TABLE 4

Male-Female Score Overlap for UNIACT Scales

Scales (Holland types in parentheses)		UNIACT	national n	orm gro	up	Percent overlap to other samples			
	Ma	Males Female		ales					
	x	SD	x	SD	Percent overlap	1 <sup>a</sup>	2 <sup>b</sup>	3 <sup>c</sup>	
Science (I)	2.20	.58	2.10	.60	93	90	95	92	
Creative Arts (A)	2.09	.51	2.26	.52	87	94	85	86	
Social Service (S)	2.34	.42	2.48	.37	85	84	85	84	
Business Contact (E)	2.16	.43	2.17	.44	99	96	98	100	
Business Detail (C)	2.01	.49	2.05	.54	<b>9</b> 7	97	97	98	
Technical (R)	1.89	.42	1.77	.44	89	82	91	87	

Note. The national norm group consists of a systematic random sample of 1,247 males and 1,693 females drawn from the 198,000 persons registering for the November 1977 ACT Assessment Program (AAP) national test date. Percent overlap is based on Dunnette's (1966) table for Tilton's (1937) measure of overlap.

they must be "less valid." Research summarized in a previous section showed that sex-balanced reporting procedures based on traditional scales were at least as valid as, and sometimes more valid than, sex-restrictive reporting procedures. Hence, sex-balanced reporting procedures provide the best comparison basis for sex-balanced (unisex) scales.

Table 5 summarizes the results of studies comparing the validity of sex-balanced score reports and sex-balanced (unisex) scales. As before, citations are provided for studies that have already appeared in the professional literature.

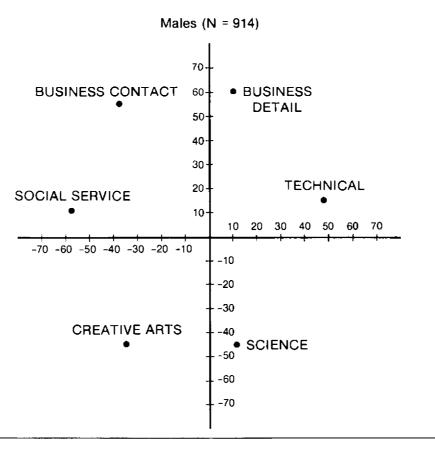
Studies 2 and 3 listed in Table 5 are the same as those described previously. Hit rates for the two studies are summarized in Table 6. The unweighted average hit rates in Table 6 indicate that results for sex-balanced scales are similar to those for sex-balanced reports. (As before, hit rates that differ by less than 5% were considered to be similar.) However, trends favor the sex-balanced reports in three of the four comparisons.

The data from Studies 2 and 3, together with the previous data summarized in Table 5, suggest a similar level of validity for unisex scales and sexbalanced score reports. As we have noted in a

<sup>&</sup>lt;sup>a</sup>Data based on cross-sectional sample of 1,851 11th graders (914 males and 937 females) attending 16 high schools in 15 states (Hanson, et al., 1977).

<sup>&</sup>lt;sup>b</sup>Data based on systematic random sample of 737 males and 852 females drawn from the 118,000 high school *seniors* registering for the October 1977 AAP national test date. Before sample selection, the population was stratified by Holland type on the basis of vocational plans.

<sup>&</sup>lt;sup>C</sup>Data based on systematic random sample of 1,297 males and 1,788 females drawn from the 127,000 persons registering for the October 1977 AAP national test date. This sample provided UNIACT norms during the 1977-78 AAP test year.



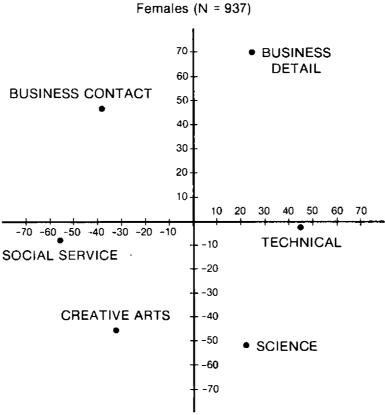


Figure 1. Plot of UNIACT theory-based factor loadings. (Source: Hanson, et al., 1977, p. 21)

TABLE 5

Summary of Validity Data for Sex-balanced Score Reports and Sex-balanced (Unisex) Scales

Study	Type of validity	Time interval	Sample; No. of males (M) & females (F)	Criterion; No. of criterion groups	Relative performance of sex-balanced reports (SBR) & sex-balanced scales (SBS)
Rayman (1976)	Construct	Concurrent	College-bound students; M=729, F=1,173	NA	SBR and SBS demonstrate similar construct validity as measures of Holland types
Hanson & Rayman (1976)	Criterion- related	Concurrent	College-bound students; M=582, <b>F</b> =878	Occ. preference; M=6, F=5 (by Holland type)	SBR and SBS discriminate among criterion groups in similar manner; hit rates similar <sup>a</sup> for males and for females
Hanson, et al. (1977); Prediger & Hanson (1978)	Construct	Concurrent	High school juniors; M=914, F=937	NA	SBR and SBS demonstrate similar construct validity as measures of Holland types
2 <sup>b</sup>	Criterion- related	Concurrent	College seniors; M=929, F=1,033	College major; M=6, F=6 (by Holland type)	SBR and SBS hit rates similar for males and for females
3 p	Criterion- related	Concurrent	College-bound students; M=737, F=852	Occ. preference; M=6, F=6 (by Holland type)	SBR and SBS hit rates similar for males and for females; differences favored SBR

Note. Sex-balanced reports (SBR) based on the application of same-sex norms to traditional interest scales are compared with scores obtained from sex-balanced (i.e., unisex) scales. All comparisons involve two interest inventories, each designed to assess Holland's six types.

recent article (Prediger and Hanson, 1978), "perfect sex balance has not been achieved with [UNIACT scales]. Indeed, there is no evidence that the vocational interests of males and females are exactly alike." But we believe that, taken as a whole, the validity data suggest "that similar interest patterns for males and females come closer to reality than the highly divergent interest patterns produced by many interest inventories." In summary, sexbalanced scales appear to provide a promising alternative for assessing basic interests, Holland types in particular.

# Concluding Comments

The field of vocational interest assessment has had 50 years of practice in constructing inventories on which males and females score differently. Given that perspective, we feel pretty good about what has been accomplished over the past four years. Fifty years of tradition in interest assessment are not easy to overcome, however.

In order to provide perspective on the reasons sexrestrictive interest assessment will be with us for a

<sup>&</sup>lt;sup>a</sup>When SBR and SBS hit rates differed by less than 5% (e.g., 46% vs. 42%), they were considered to be similar.

<sup>&</sup>lt;sup>b</sup>See descriptions of Studies 2 and 3 in this paper.

TABLE 6

Criterion Group Hit Rates for Sex-balanced (Unisex) Scales and Sex-balanced Score Reports for Holland Types

				Hit rates	s (in %)	
Criterion group	Samp	le size		lanced ) scales <sup>a</sup>	Sex-balanced reports b	
by Holland type	M	F	M	F	M	F
			Study 2			
Investigative	323	348	53	55	46	50
Artistic	148	188	63	61	79	57
Social	151	182	27	32	30	24
Enterprising	121	121	56	46	62	43
Conventional	105	118	42	61	47	55
Realistic	81	76	33	22	<b>3</b> 3	47
Unweighted						
average hit rate			46	46	50	46
			Study 3			
Investigative	187	181	43	22	40	32
Artistic	142	187	45	48	55	41
Social	76	132	29	29	41	26
Enterprising	124	145	41	31	37	34
Conventional	101	132	64	51	69	62
Realistic	107	75	41	23	36	35
Unweighted average hit rate			44	34	46	38

Note. All comparisons involve two interest inventories, each designed to assess Holland's six types.

<sup>&</sup>lt;sup>a</sup>Scales consist of items for which males and females give similar responses. Reports are based on standard scores derived from combined-sex norms.

<sup>&</sup>lt;sup>b</sup>Reports are based on standard scores derived from same-sex norms.

long time, I have assembled a list of 11 propositions that have been made in the professional literature—though in a more seductive manner. The propositions are uncontaminated by the results of research. Nevertheless, we believe they deserve serious (but not too serious) attention.

- Once sex-role socialization has taken hold, a counselee's vocational options are restricted for life. Corollary: Because sex-restrictive interest scores simply reflect the effects of socialization, the only valid way to eliminate sexrestrictive scores is to revise society. Revising society is easier than revising interest inventories. But, for all of your counselees, it's already too late.
- Vocational interest inventories must reflect sexrole stereotypes so that we can know when those stereotypes change. Corollary: Vocational counseling must also reflect sex-role stereotypes. Counselees can come back several years later to see if their stereotypes have changed.
- 3. The items used on vocational interest inventories (e.g., would you like to operate a power shovel? Repair a hot rod? Drill soldiers? Tend babies?) are "gender neutral." Hence, any differences in the interest scores of males and females simply reflect a FACT OF LIFE. Corollary: Interest inventories on which males and females receive similar scores are not just invalid, they mess with Mother Nature.
- 4. Interest inventories that suggest similar vocational options to males and females are difficult to reconcile with current theories of vocational development. (Current theories leave no doubt that males and females are destined for different occupations. Mother Nature approves.)
- 5. If one develops an interest inventory on which males and females receive similar scores, one must do the same for Bohemians, Unitarians, Middle Americans, and card-carrying Democrats—regardless of whether bias exists for any of these groups. The result will be an inventory with no more than two or three items. It will make everyone appear equal.
- The correct way to validate an interest inventory is to see how well it predicts which occupations counselees will eventually enter or prefer.

Counselors make heavy use of such predictions and counselees find them simply amazing.

- 7. Sex-restrictive inventories are "more valid" than non-sex-restrictive inventories. (In predicting future occupational entry or preference, they are almost as valid as predictions based on a counselee's stated occupational preference.)
- 8. Interest inventories should suggest occupations that parallel the traditional employment distributions and stated preferences of males and females. Corollary: An interest inventory that suggests nontraditional occupations to a counselee not only messes with Mother Nature, it prevents quick closure of the case.
- 9. If the "effects" of an interest inventory on males and females are "similar" (e.g., if both males and females explore the occupations suggested by their scores), then the inventory is SEX FAIR—even if the suggested occupations are highly sex stereotypic. Corollary: Counselors who produce "similar effects" on male and female counselees are also SEX FAIR. However, claims regarding male chauvinist counselors are ambiguous.
- 10. There are numerous purposes for using interest inventories in vocational counseling—for example, to enhance self-knowledge and identify career alternatives; to enhance self-knowledge and identify career alternatives for exploration; to (etc.) and compare career alternatives with current expressed choices; to (etc.) for college sophomores, disco dancers, clone donors, and near-sighted left fielders. The number of purposes is SO GREAT that issues of sex bias can never be resolved by empirical research or scientific reason.
- 11. There are numerous definitions of sex bias in interest inventories. Until everyone agrees on a single definition, sex bias can't even be identified, much less eliminated. Corollary: Efforts to eliminate racism will also be useless until everyone agrees on a single definition.

For all of the above reasons, sex-restrictive interest inventories will be with us for a long time. Yet, millions of vocational interest inventories are used, year in and year out, by vocational counselors and others in the helping professions. Research has shown that both sex-restrictive and sex-balanced

interest reports produce increased exploration of the vocational options that are suggested (Prediger & Hanson, 1976; Prediger & Noeth, in press). Each year, many persons make vocational plans grounded, at least in part, on sex-restrictive reports of basic vocational interests. It is difficult to imagine a clearer example of a problem that needs to be and can be addressed.

#### **Occupational Interest Scales**

Books on measurement typically cite the Strong Vocational Interest Blank (SVIB), or its successor, the Strong-Campbell Interest Inventory (SCII) (Campbell, 1977), as a model for occupational interest scale development. In a recent survey of testing practices at university and college counseling centers, 94% of the respondents (269 of 284 centers) stated that they used the SCII in counseling students (Sell & Torres, Note 1). One-half of the respondents (142 of 284) indicated that the majority of their clients completed the SCII. No other psychological test approached this degree of use.

In this section of the report, sex restrictiveness in occupational interest scales is examined, and strategies for reducing sex restrictiveness are discussed. Because of its widespread recognition and use, major attention is given to the SCII. The discussion extends to occupational interest scales on other instruments, however, and the use of basic interest scales to clarify the interpretation of occupational scales is considered.

### The Problem of Sex Restrictiveness

Different sets of occupational scales. The SCII and two alternative interest inventories, the Career Assessment Inventory (CAI) (Johansson, 1976) and the Kuder Occupational Interest Survey (KOIS), Form DD, (Kuder, 1975), contain separate sets of occupational scales for men and women. Of the 124 occupational scales on the SCII, there are 30 scales for men and 20 scales for women that are not matched by scales for the opposite sex. Scales such as Highway Patrol Officer, Skilled Crafts, and Investment Fund Manager have been developed only for males; scales such as Dental Hygienist, Secretary, and Beautician exist only for females. Similarly, 22 of the 42 scales on the CAI have been constructed for just one sex (14 for men and 8 for women), and 80 of the 114 scales on the KOIS pertain to only one sex (60 for men and 20 for women).

The predominant types of interests represented by the female and male occupational scales on each of the three inventories are shown in Table 7. For women, there are proportionately fewer scales representing Holland's (1973) realistic or enterprising types of occupations; for men, there are proportionately fewer scales representing conventional types of occupations. These differences are pronounced for all three inventories. In addition, for the CAI and KOIS, there is a much larger percentage of scales indicating social types of occupations for women than for men.

The differences in the types of interests represented by the occupational scales for men and women reflect the differences in the employment patterns of men and women. The use of these scales in their present form serves to perpetuate the status quo. As an extreme example, there are no female occupational scales on the KOIS in Holland's realistic category, while more than one-quarter of the male scales fall in this category. It would be difficult for a female to show interests compatible with occupations in the realistic category on the present form of this instrument.

Prediger (1977) has argued that interest inventories should be used to suggest possible occupations for exploratory purposes rather than to predict future occupational membership. If the emphasis is removed from predictive accuracy to career exploration, the rationale for establishing different types of scales for men and women is less valid. To encourage career exploration, the full range of career possibilities should be adequately represented for both men and women. As described in a previous section entitled "Validation Models," new types of studies to determine the validity of interest inventories for enhancing career exploration would be required.

Differences in mean scores. Sex restrictiveness in the interest inventories is also shown in the high scores obtained by men and by women. When only

TABLE 7

Types of Interests Represented by Female and Male Occupational Scales on Three Interest Inventories

		Strong-C Interest I	•			Career As		nt Kuder Occupation Survey, For				
Tuno of	Femal	e scales	Male	scales	Femal	e scales	Male	scales	Femal	Female scales		scales
Type of interest <sup>a</sup>	N	%	N	%	N	%	N	%	N	%	N	%
Realistic	5	8.8	14	20.9	1	5.6	9	37.5	0	0	20	26.0
Investigative	19	33.3	14	20.9	1	5.6	1	4.2	8	21.6	20	26.0
Artistic	10	17.5	9	13.4	3	16.7	3	12.5	2	5.4	5	6.5
Social	8	14.0	12	17.9	3	16.7	2	8.3	18	48.6	13	16.9
Enterprising	5	8.8	15	22.4	3	16.7	6	25.0	3	8.1	15	19.5
Conventional	10	17.5	3	4.5	7	38.9	3	12.5	6	16.2	4	5.2
Total	57	99.9	67	100.0	18	100.2	24	100.0	37	99.9	77	100.1

<sup>&</sup>lt;sup>a</sup>The Occupational scales have been classified according to predominant type of interest by Campbell (1977, p. 16), Johansson (1976, p. 65), and Diamond (1975, pp. 4-5). Two scales on the Kuder Occupational Interest Survey not classified by

Diamond—i.e., the Forester (male) and Home Demonstration Agent (female) Scales—were placed in the realistic and enterprising categories, respectively.

parallel scales on the SCII are considered, men score high ( $\geq$ 35) on realistic and investigative occupational scales (Engineer, Computer Programmer, and College Professor), while women score high on artistic and social occupational scales (Musician, Physical Therapist, Elementary Teacher, and Recreation Leader) (Campbell, 1977, p. 74).

Among the ten like-named male and female scales on the CAI, females score highest (> 30) on three female occupational scales classified as conventional (Accountant, Executive Housekeeper, and Food Service Manager), while men score highest on two male scales in the enterprising category (Buyer/Merchandiser and Food Service Manager) (Johansson, 1976, p. 68). Comparable data for the KOIS could not be located, but there appears to be a tendency for men to score highest on realistic. investigative, and enterprising occupational scales and for women to score highest on social, conventional, and artistic scales (Kuder, 1975). These data indicate that different career options will be suggested to men and women even when identical sets of scales are presented to both sexes.

### Alternative Methods of Scale Construction

It might be possible to reduce sex restrictiveness by introducing new occupational scales that would offer a more balanced list of career options to men and women. Researchers have considered at least four different types of new scales: (a) cross-sex scales, (b) combined-sex scales, (c) cluster scales, and (d) sex-balanced scales.

Cross-sex scales. One way of exposing men and women to a broader variety of occupational possibilities is to use both the female and male scales with the same client, regardless of sex (Schlossberg & Goodman, 1972). This procedure has been adopted recently for all three inventories cited above.

Ironically, the use of the cross-sex scales may limit career exploration rather than expand it. In general, individuals taking the SCII obtain elevated scores on the cross-sex scales that represent occupations traditional for their sex and depressed scores on the cross-sex scales that represent nontraditional occupations (Johnson, 1977; Lunneborg, 1975). For example, women score relatively high on "artistic" male occupational scales and relatively low on

"realistic" male occupational scales. Men obtain the opposite results on the female scales. The use of the cross-sex scales tends to reinforce sexual stereotypes and to hinder the consideration of nontraditional occupations. Johansson (1976) notes that similar results occur when the cross-sex scales on the CAI are used.

A somewhat different result occurs when the crosssex scales on the KOIS are employed. Because of the scale construction procedures, most individuals will receive lower scores on the cross-sex scales than on the same-sex scales (Diamond, 1974). If the scores are interpreted without reference to the sex group, use of the cross-sex scales will suggest relatively few new occupations. (An alternative approach to KOIS interpretation is described later.)

Combined-sex scales. Several attempts have been made to construct occupational scales based upon combined samples of men and women. Years ago, Strong (1943, pp. 568-576) noted that men's and women's scales could be "combined in some cases" (e.g., Artist), but not in others (e.g., Lawyer).

More recently, Webber and Harmon (Note 2) found that scales based upon combined-sex samples of veterinarians and life insurance agents were more effective in identifying female veterinarians and female life insurance agents than were scales based upon female samples. In contrast, the male Veterinarian and Life Insurance Agent Scales were more successful in identifying the male members of these occupations than were the combined-sex scales.

Hanson (1976) obtained similar results in experimenting with different versions of a scale for sociologists. The combined-sex scale was slightly more effective than the female scale in differentiating female sociologists from other groups of females. The male scale proved to be more accurate in identifying male sociologists than did the combined-sex scales.

In a study with the CAI, Johansson (1976) found that single-sex scales more clearly differentiated between male or female interior decorators and the general reference samples than did a combined-sex scale. He concluded that separate-sex scales "produce the best validity" (p. 67).

No simple conclusion can be drawn from these studies. The possibility of creating combined-sex scales deserves further study. However, different types of scales may be needed for men and women or for different occupations.

Cluster scales. Cluster scales have been developed for the version of the KOIS that is included in the Career Development Inventory (CDI) (Borgen, 1978; Diamond, 1975). With the cluster scales, the scores on the male and female scales first are "equated for sex differences" so that the same norms may be used for both sexes. Scores are then averaged for scales within each of the six Holland categories. A further distinction is made in terms of the level of the occupation (early entry, delayed entry, or late entry) so that 18 occupational cells or categories are created. The authors maintain that the clusters smooth out small sex differences that may appear in the specific male and female occupational scales.

This type of scale is reminiscent of the occupational group scales previously used with the SVIB, except that the same scales are used with males and females. It should be noted that, when the cluster scales are used, large differences still exist in the percentages of high scores obtained by men and women in the various categories. For example, 64% of the males in a study conducted with the CDI obtained high scores in the realistic (technical/mechanical/skilled), delayed-entry category, while only 13% of the females obtained a high score in this category (Diamond, 1975).

Sex-balanced scales. Finally, it may be possible to eliminate sex restrictiveness in the occupational scales by using only those interest items that are preferred equally by men and women. As noted in the discussion of unisex scales in the first part of this report, scales based on such items are "sex balanced" in that males and females obtain approximately equal scores. In contrast to the traditional scales, combined-sex norms may be used with sexbalanced scales; males and females will still obtain approximately equal scores.

The greatest problem in constructing sex-balanced interest scales for the SCII, CAI, or KOIS is the lack of sex-balanced items. Large differences in the item preferences of men and women exist for approximately one-half of the items on the SCII and CAI (Campbell, 1977; Johansson, 1976). Thus, unless new interest items are constructed, scale lengths must be reduced substantially or items with less validity must be used if sex-balanced interest scales are to be developed.

Several studies have investigated the relative validity of sex-balanced and traditional occupational scales (Hansen, 1976; Webber & Harmon, Note 2: Johnson, 1978). In each of the studies, items that were not sex balanced were eliminated from the traditional scales in order to form sexbalanced scales. As a result, the number of items on each of the sex-balanced scales either was reduced substantially or was maintained by adding items that were less effective than the original items in differentiating between occupational groups. In most cases, the traditional scales were more valid than the sex-balanced scales; however, the differences were fairly small. None of the differences in the amount of overlap between men or women in the occupation and men- or women-in-general exceeded five percentage points. Comparisons between sex-balanced and traditional occupational scales with an equal number of items of comparable validity have yet to be reported. Whether it is more difficult to write sex-balanced items for occupational scales than for basic interest scales remains to be seen. As noted above, approximately one-half of the current items written for the SCII and CAI are sex balanced.

In any case, research with the MMPI shows that shortened scales may serve many of the purposes of the original scales with relatively little loss in reliability and validity (Kincannon, 1968; Freeman, O'Leary, & Calsyn, 1977). This line of research suggests that an abbreviated version of the occupational scales based only on sex-balanced items may be a possibility. Because of their reduced length, these scales would not be as reliable over long time periods (Johnson, 1978). In addition, these scales would probably not be as accurate in predicting occupational membership some years in the future (should that be a counselor's goal), but they should prove to be helpful in expanding the career options of men and women.

Alternative Methods of Interpreting Scores on Existing Scales

Despite attempts to construct new scales, the existing occupational scales will probably continue to be used for some years. Most of the new scales developed for the SCII have followed the traditional design (Aburto, Note 3; Hansen, Notes 4 & 5; Larkin, Note 6; Stocco, Note 7). For this reason, it is important for counselors to consider how the existing scales can be interpreted in a way that will reduce sex restrictiveness.

Using separate sex norms. Increased normative data are needed in interpreting the scores of men and women on the occupational scales. If possible, this information should be provided on the interest profile for easy reference. The range of scores for the middle third of men- or women-in-general, depending on the type of scale, is shown on the CAI profile for the occupational scales. This type of information should be portrayed for both sexes on each scale. Normative data for either sex is lacking on the SCII and the KOIS profiles. Although it will require some ingenuity to design a profile to permit addition of these data, the introduction of this type of information on the computer printout forms should be relatively easy.

Additional data to help in interpreting the scores of males and females on the cross-sex scales are shown in Table 8. This table, which shows the first, second, and third quartiles separately for men and women on each scale, is derived from the scores of 1,134 male and 1,044 female freshmen at the University of Wisconsin–Madison. This sample, tested with a prepublication version of the SCII in 1973, represented 50.1% of the entering freshmen males and 51.7% of the entering freshmen females. These students closely resembled the other students in academic achievement as measured by the College Qualifications Tests. Few SCII norms based on college students have been published.

The data in Table 8 show that the scores on the occupational scales may vary considerably for men and women. For example, a score of 24 on the male Farmer Scale is average (50%ile) for men but above average (75%ile) for women. A score of 36 on the male Dietitian Scale is average for women but above average for men. If the sex norms are not taken into account, "farmer" would more often be suggested to males as a career option, while "dietitian" would more often be suggested to females. These suggestions would reinforce sexual stereotypes.

Because of sex-role conditioning, the scores obtained on interest scales convey different meanings for men and women. For example, successful female science majors do not express as many mechanical interests as do successful males in this major (Goldman, Kaplan, & Platt, 1973). Separate sex norms are needed to take into account discrepancies in social conditioning that may be inhibiting the endorsement of certain types of interests.

Separate sex norms are needed on the SCII and CAI where the scores are relatively high on some cross-sex scales and relatively low on others. For the KOIS, the scores on the cross-sex scales apparently are lower in almost all cases. In this situation, interpreting the rank order of each set of scales for each counselee may suffice (Tittle & Denker, 1977). Zytowski and Laing (1978) found that cross-sex scales on the KOIS were as valid as same-sex scales in predicting occupational membership when the rank-order of the occupational scores was considered separately for the male and female scales.

Using relationships between occupational and basic interest scales. In addition to knowing the relative magnitude of their scores on the cross-sex scales, clients should be given information on the nature of these scores. As one step in this direction, correlation coefficients between occupational scales and basic interest scales should be helpful in identifying underlying interest patterns. Correlations for the SCII, based on data from the Wisconsin sample described above, are reported in the appendix.

The correlations show, for example, that high scores on the Farmer Scale are most closely associated with low scores on the Writing (r = -.81), Music (r = -.66), and Art (r = -.60) Basic Interest Scales and, to a lesser degree, with high scores on the Mechanical (r = .28). Athletics (r = .24), and Agriculture (r = .22) Scales. These results suggest that a woman who states a preference for the SCII artistic activities, as most women do, would probably obtain a low score on the Farmer Scale even if she liked agricultural and mechanical activities. This low score may be more accurately evaluated by using female norms and by noting the Basic Interest Scales that may be contributing most to the score. If the Basic Interest Scales are not relevant for the expression of that occupation, they should be given less consideration (Johnson & Johansson, 1972).

High scores on the male Dietitian Scale, on the other hand, were most closely related to high scores on the Domestic Arts (r = .67), Office Practices (r = .66), and Medical Service (r = .63) Basic Interest Scales. These scales measure activities usually preferred by women. To the extent that individuals respond to interest items in a stereotypic fashion, the scores for females will be artifically inflated on the male Dietitian Scale. The

TABLE 8

Quartile Scores of Male and Female Freshmen at the University of Wisconsin-Madison on the SCII Occupational Scales

Occupational scale		·	Males			Females		
Title	Sex <sup>a</sup>	25%ile	50%ile	75%ile	25%ile	50%ile	75%ile	
armer	m	15	24	33	11	17	24	
nstrument Assembler /ocational Agriculture	f	23	32	40	17	25	34	
Teacher	m	6	16	25	-1	6	15	
Dietitian	m	23	29	36	29	36	42	
olice Officer	m	19	29	39	11	19	28	
lighway Patrol Officer	m	16	26	36	5	12	21	
Army Officer Physical Education	f	36	43	49	22	28	37	
Teacher	f	23	32	42	19	29	39	
Skilled Crafts	m	17	27	37	9	16	23	
orester	m	20	29	37	15	23	31	
Radiologic Tech.	f	26	36	43	24	34	45	
Merchant Marine Officer	m	33	3 <b>9</b>	46	27	32	37	
lavy Officer	m	18	28	38	6	15	25	
Nurse, R.N.	m	19	28	36	28	36	45	
/eterinarian	m	15	23	33	18	28	36	
Cartographer	m	31	40	49	21	27	37	
Army Officer	m	18	26	46	10	17	24	
Air Force Officer	m	18	27	37	11	17	25	
Occupational Therapist	f	21	30	38	27	35	43	
Engineer	f	25	36	46	13	21	31	
Engineer	m	23	33	42	16	22	31	
Chemist	f	13	25	38	-1	9	22	
Physical Scientist	m	16	26	36	17	26	34	
Medical Technologist	f	26	36	47	16	26	37	
Pharmacist	f	24	<b>33</b>	44	17	27	40	
Dentist	f	25	35	42	17	26	35	
Dentist	m	23	30	39	23	32	41	
Dental Hygienist	f	21	28	37	24	34	43	
Physical Therapist	f	28	37	47	26	36	47	
Physician	m	18	28	39	18	27	37	
Math-Sci <b>e</b> nce Teacher	m	24	34	45	19	28	38	
Math-Science Teacher	f	31	39	46	22	29	36	
Dietitian	f	19	27	35	23	31	39	
Medical Technologist	m	13	24	36	9	20	33	
Optometrist	m	20	30	39	15	26	35	
							(Contin	

23

TABLE 8—Continued

Occupational scale		Males			Females		
Title	Sex <sup>a</sup>	25%ile	50%ile	75%ile	25%ile	50%ile	75%ile
Computer Programmer	f	28	37	47	15	26	35
Computer Programmer	m	23	34	45	17	25	33
Mathematician	f	23	31	39	12	20	29
Mathematician	m	17	25	35	21	29	36
Physicist	f	15	25	38	-3	7	19
Biologist	m	17	26	35	26	34	41
Veterinarian	f	28	36	43	20	29	38
Optometrist	f	29	38	45	18	25	34
Physician	f	28	37	45	18	27	37
Social Scientist	m	22	31	40	31	38	46
Speech Pathologist	f	25	33	41	24	31	40
Speech Pathologist	m	19	30	40	32	40	48
College Professor	f	35	42	48	29	36	42
College Professor	m	30	39	46	37	43	50
Psychologist	f	20	28	35	14	24	33
Psychologist	m	19	28	38	26	34	43
_anguage Interpreter	f	25	31	40	22	31	41
Architect	m	10	19	28	12	23	31
Advertiser	f	30	36	42	24	31	39
Artist	f	20	29	39	19	29	40
Artist	m	18	28	39	30	38	48
Art Teacher	f	0	9	21	8	21	31
Photographer	m	17	26	37	26	36	46
Musician	f	15	25	35	29	37	45
Musician	m	29	39	48	40	48	56
Entertainer	f	19	26	34	28	36	43
nterior Decorator	f	2	10	19	4	15	25
nterior Decorator	m	17	23	30	31	38	43
Advertiser	m	19	28	37	26	35	44
anguage Teacher	f	7	16	28	23	32	40
ibrarian	f	19	27	36	17	29	39
ibrarian	m	13	22	32	27	34	43
Reporter	f	22	31	39	22	32	41
Reporter	m	20	29	39	31	39	47
English Teacher	f	10	20	32	20	31	41
English Teacher	m	18	27	39	30	39	47
Nurse, R.N.	f	17	25	32	23	32	43
Physical Therapist	m	21	29	37	19	27	36
Nurse, Licensed Practical	m	24	29	35	35	41	47
Social Worker	f	11	21	33	16	26	37

TABLE 8—Continued

Occupational scale		' Males			Females			
Title	Sex <sup>a</sup>	25%ile	50%ile	75%ile	25%ile	50%ile	75%ile	
Social Worker	m	9	20	31	19	28	37	
Priest	m	14	24	34	26	34	42	
Director, Christian Ed.	f	2	9	20	11	21	31	
WCA Staff	f	21	29	38	24	33	41	
Minister	m	12	20	30	20	28	36	
lementary Teacher	m	17	27	36	· 27	35	42	
lementary Teacher	f	12	19	26	20	29	36	
School Superintendent	m	13	21	30	15	23	31	
Public Administrator	m	21	29	3 <b>9</b>	21	27	34	
Suidance Counselor	m	15	22	30	18	26	34	
Recreation Leader	m	13	23	33	15	23	33	
Suidance Counselor	f	9	19	30	17	26	36	
Social Science Teacher	f	19	30	40	22	30	38	
Social Science Teacher	m	19	28	36	22	30	37	
Personnel Director	m	20	27	36	21	27	33	
Department Store								
Manager Home Economics	m	14	21	29	12	18	27	
Teacher	f	-5	2	9	9	20	32	
Stewardess	f	16	23	31	22	31	40	
Chamber of Commerce								
Executive :	m	21	28	37	21	27	35	
Sales Manager	m	.12	19	27	10	16	23	
ife Insurance Agent	m	10	17	25	9	16	23	
ife Insurance Agent	f	17	24	33	15	21	28	
awyer	f	29	36	44	18	28	37	
awyer	m	18	28	38	23	31	39	
Computer Sales	m	9	17	25	5	12	20	
nvestment Fund Manager	m	22	29	37	22	28	34	
Pharmacist	m	17	26	36	15	25	34	
Buyer	f	16	23	30	16	20	27	
Buyer	m	9	16	27	7	16	26	
Credit Manager	m	15	25	34	12	21	29	
Funeral Director	m	; 18	26	35	18	26	34	
Realtor	m	19	26	35	14	19	26	
Agribusiness Manager	m	. 7	17	27	2	10	19	
Purchasing Agent	m	19	28	37	14	20	28	
Chiropractor	m	25	32	40	26	33	38	
•		•					(Contin	

25

**TABLE 8**—Continued

Occupational scale		Males			Females		
Title	Sex <sup>a</sup>	25%ile	50%ile	75%ile	25%ile	50%ile	75%ile
Accountant	m	6	15	25	-4	4	14
Banker	f	22	30	36	17	22	30
Banker	m	15	23	32	10	16	23
Credit Manager	f	20	28	37	13	21	29
Department Store Sales	f	9	16	23	12	19	27
Business Education							
Teacher	f	9	14	21	8	14	21
Business Education							
Teacher	m	15	24	32	14	22	30
Executive Housekeeper	f	11	18	25	13	22	29
Accountant	f	20	29	37	9	17	24
Secretary	f	19	25	32	21	28	3 <b>6</b>
Dental Assistant	f	18	25	31	17	, 27	38
Nurse, Licensed Practical	f	13	18	25	14	<sup>23</sup>	32
Beautician	f	23	29	36	24	31	37

Note. The data are based on 1,134 male and 1,044 female first-year students at the University of Wisconsin-Madison.

elevated scores for females may be taken into account by the use of separate sex norms and by a better understanding of the types of responses that are producing these scores.

Several authorities have urged counselors to place greater emphasis in career planning on homogeneous interest scales such as the SCII General Occupational Theme Scales or the SCII Basic Interest Scales (Cole & Hanson, 1971; Harmon, 1975). Scores on these scales are often much more revealing than scores on the occupational scales. For example, female and male radiologic technologists have similar interest patterns on the SCII Basic Interest Scales when the appropriate sex norms are used (Stocco, Note 7). Both sexes score above average on the Medical Service and Medical Science Scales and within the average range on the other scales. In contrast, female and male farmers

produce somewhat different interest patterns (Hansen, Note 4). Although both sexes obtain high scores on the Agriculture and Nature Scales, male farmers also score relatively high on the Mechanical Activities Scale while female farmers score relatively high on the Office Practices, Sales, Domestic Arts, and Religious Activities Scales. These differences reflect differences in the work roles of many farm husbands and wives. Knowledge of the basic interest patterns helps clarify the differences between men and women within the occupation at the present time without necessarily suggesting that there is a preferred interest pattern for each sex.

As a further example of how basic interest scales aid an interpretation, consider the scores on the SCII Basic Interest Scales and the Psychologist Occupational Scales shown in Table 9. The female counselee who obtained these scores had a 33 on

af = female scale; m = male scale.

TABLE 9

Scores on the SCII Basic Interest Scales and Psychologist Scales for a Female Counselee

SCII Scale	Score	Interpretation <sup>a</sup>
Basic Interest Scale		
Agriculture	46	Below average
Nature	54	Average
Adventure	34	Below average
Military Activities	41	Below average
Mechanical Activities	47	Below average
Science	62	Above average
Mathematics	65	Above average
Medical Science	<b>6</b> 3	Above average
Medical Service	65	Above average
Music/Drama	64	Above average
Art	52	Average
Writing	56	Average
Teaching	55	Average
Social Service	57	Average
Athletics	39	Below average
Domestic Arts	58	Average
Religious Activities	56	Average
Public Speaking	45	Below average
Law/Politics	40	Below average
Merchandising	54	Average
Sales	52	Average
Business Management	55	Average
Office Practices	62	Above average
Occupational Scale		
Psychologist (male)	52	Above average
Psychologist (female)	33	Average

<sup>&</sup>lt;sup>a</sup>Based on female norms (Above average =>75%ile; Average = 25%ile-75%ile; Below average =<25%ile).

the female Psychologist Scale and a 52 on the male Concluding Comments Psychologist Scale. Even when the female norms for the male scale are used, her score still is considerably higher on the male scale than it is on the female scale.

The counselee, a 26-year-old college graduate with honors in a health-related field, wished to consider various academic or career options, including psychology. Should her enthusiasm for psychology be dampened because of her relatively low score on the female Psychologist Scale? Inspection of the Basic Interest Scales indicated that she shared many of the interests of psychologists (e.g., see scores for the Medical Science, Science, Mathematics, and Medical Service Scales). She differed from female psychologists in that she also scored relatively high on the Office Practices, Religious Activities, and Domestic Arts Scales. These latter interests, which are not scored as frequently on the male scale as on the female scale, are weighted negatively on the female scale (Johnson, 1974; Johnson & Campbell, 1974). Her score on the male Psychologist Scale was also higher than her score on the female scale because the male scale gives more weight to her interests in music and writing. The latter activities differentiate male psychologists from other males more effectively than they differentiate female psychologists from other females.

Because of the complex nature of the scores on the male and female Psychologist Scales, it is more helpful to discuss the actual interest patterns of psychologists than to emphasize a single score. In this case, the client decided that the incongruence between her interest pattern and that of female psychologists was not critical. She felt that she could consider psychology as a major or a career and still express her interests in office practices. religious activities, and domestic arts within the life style she established.

The male and female occupational interest scales on the SCII, CAI, and KOIS are sex restrictive inasmuch as they suggest different career options for males and females. Different types of unisex scales have been constructed in an attempt to reduce the sex restrictiveness of these three interest inventories. The use of sex-balanced interest items in constructing unisex scales appears promising. However, approximately one-half of the items on the interest inventories cannot be used for this purpose. Whether these items can be replaced with new sex-balanced items remains to be seen.

The possibility of developing a short form of the occupational scales based only on sex-balanced interest items is worth consideration. Although such a short form may provide less accurate predictions of future choice than the regular form, its construct validity may be sufficient for use in stimulating career exploration. As previously noted, prediction of future vocational choice is seldom a goal in counselina.

From a practical standpoint, it may be more helpful to improve the methods of interpreting the existing occupational scales than to attempt to create a large number of unisex scales within the near future. There is a need for appropriate sex norms and for additional data to clarify the nature of the interest patterns inherent in the occupational scores. Some of these data are provided in this report. Interpretive aids of this sort should help clients take into account the influence of sex-role conditioning on occupational interest scores.

Sex restrictiveness in interpretations of the SCII Occupational Scales may also be reduced by placing increased emphasis on the Basic Interest Scales or the General Occupational Theme Scales. When separate-sex norms are used, the same types of career options are suggested to males and females by means of these scales.

#### Reference Notes

- Sell, J.M., & Torres, R. Test usage in college and university counseling centers. Unpublished manuscript, Counseling and Testing Center, Southern Methodist University, 1977.
- Webber, P. L., & Harmon, L. W. A concurrent validity study of single-sex and combined-sex scales. Paper presented at the meeting of the American Personnel and Guidance Association, New Orleans, April 1974.
- 3. Aburto, R. What do newcomer leaders of the church (F-ministers) look like? Paper presented at the meeting of the American Personnel and Guidance Association, Washington, D.C., March 1978.
- Hansen, J. C. Farm wives: Will their self-image ever change? Paper presented at the meeting of the American Personnel and Guidance Association, Washington, D.C., March 1978.

- Hansen, J. C. The Strong-Campbell Interest Inventory: A glimpse of the 1980's. Paper presented at the meeting of the American Personnel and Guidance Association, Washington, D.C., March 1978. (a)
- 6. Larkin, K. The vocational interests of male dietitians. Paper presented at the meeting of the American Personnel and Guidance Association, Washington, D.C., March 1978.
- 7. Stocco, J. L. A male radiologic technologist scale for the SCII. Paper presented at the meeting of the American Personnel and Guidance Association, Washington, D.C., March 1978.

	•			
	•			
ì.				
<b>`</b>				
<b>.</b>				
<b>.</b>				
<b>.</b>				
: :				

#### **REFERENCES**

- AMEG Commission on Sex Bias in Measurement. AMEG Commission report on sex bias in interest measurement. *Measurement and Evaluation in Guidance*, 1973, 6, 171-177.
- The American College Testing Program. Vocational Interest, Experience, and Skill Assessment: User's handbook. Boston: Houghton Mifflin, 1976.
- Anastasi, A. *Psychological testing* (4th ed.). New York: Macmillan, 1976.
- Berdie, R. F. Vocational choice studies. *Personnel and Guidance Journal*, 1970, 48, 604-605.
- Bingham, R. P., & Walsh, W. B. Concurrent validity of Holland's theory for college-degreed black women. *Journal of Vocational Behavior*, 1978, 13, 242-250.
- Borgen, F. H. Test review: Career Development Program. *Measurement and Evaluation in Guidance*, 1978, 10, 244-247.
- Boyd, V. S. Neutralizing sexist titles in Holland's Self-Directed Search: What difference does it make? *Journal of Vocational Behavior*, 1976, 9, 191-199.
- Campbell, D. P. Manual for the Strong-Campbell Interest Inventory (2nd ed.). Stanford, Calif.: Stanford University Press, 1977.
- Cole, N. S., & Hanson, G. R. An analysis of the structure of vocational interests. *Journal of Counseling Psychology*, 1971, 18, 478-486.
- Cole, N. S., & Hanson, G. R. Impact of interest inventories on career choice. In E. E. Diamond (Ed.), Issues of sex bias and sex fairness in career interest measurement (National Institute of Education Report). Washington, D.C.: U.S. Government Printing Office, 1975.
- Diamond, E. E. Handling sex bias in Kuder Occupational Interest Survey, Form DD: Some questions and answers. Chicago: Science Research Associates, 1974.

- Diamond, E. E. Career Development Program technical supplement: Career Development Inventory. Chicago: Science Research Associates, 1975.
- Dunnette, M. D. Personnel selection and placement. Belmont, Calif.: Wadsworth, 1966.
- Equal Employment Opportunity Commission. Guidelines on employment selection procedures. Federal Register, 1 August 1970, 35, 12.333.
- Fishburne, F. J., Jr., & Walsh, W. B. Concurrent validity of Holland's theory for non-college-degreed workers. *Journal of Vocational Behavior*, 1976, 8, 77-84.
- Freeman, C. W., O'Leary, M. R., & Calsyn, D. Application of the Faschingbauer Abbreviated MMPI with alcoholic patients. *Journal of Clinical Psychology*, 1977, 33, 303-306.
- Goldman, R. D., Kaplan, R. M., & Platt, B. B. Sex differences in the relationship of attitudes toward technology to choice of field of study. *Journal of Counseling Psychology*, 1973, 20, 412-418.
- Gottfredson, G. D. A note on sexist wording in interest measurement. *Measurement and Evaluation in Guidance*, 1976, 8, 221-223.
- Gottfredson, G. D., & Holland, J. L. Some normative self-report data on activities, competencies, occupational preferences, and ability ratings for high school and college students, and employed men and women. *JSAS Catalog of Selected Documents in Psychology*, 1975, 5, 192. (Ms. No. 859) (a)
- Gottfredson, G. D., & Holland, J. L. Vocational choices of men and women: A comparison of predictors from the Self-Directed Search. *Journal of Counseling Psychology*, 1975, 22, 28-34. (b)
- Gottfredson, G. D., Holland, J. L., & Gottfredson, L. S. The relation of vocational aspirations and assessments to employment reality. *Journal of Vocational Behavior*, 1975, 7, 135-148.

- Hansen, J. C. Exploring new directions for Strong-Campbell Interest Inventory occupational scale construction. *Journal of Vocational Behavior*, 1976. 9, 147-160.
- Hanson, G. R. Assessing the interests of college youth: Summary of research and applications (ACT Research Report No. 67). Iowa City, Iowa: The American College Testing Program, 1974.
- Hanson, G. R., Prediger, D. J., & Schussel, R. H. Development and validation of sex-balanced interest inventory scales (ACT Research Report No. 78). Iowa City, Iowa: The American College Testing Program, 1977.
- Hanson, G. R., & Rayman, J. Validity of sex-balanced interest inventory scales. *Journal of Vocational Behavior*, 1976, 9, 279-291.
- Harmon, L. W. Technical aspects: Problems of scale development, norms, item differences by sex, and the rate of change in occupational group characteristics. In E. E. Diamond (Ed.), Issues of sex bias and sex fairness in career interest measurement (National Institute of Education Report). Washington, D.C.: U.S. Government Printing Office, 1975.
- Harrington, T. J., & O'Shea, A. J. Manual for the Harrington/O'Shea Systems for Career Decision-making. Needham, Mass.: Career Planning Associates, 1976.
- Holland, J. L. Professional manual for the Self-Directed Search. Palo Alto, Calif.: Consulting Psychologists Press, 1972.
- Holland, J. L. Making vocational choices. A theory of careers. Englewood Cliffs, N.J.: Prentice-Hall, 1973.
- Holland, J. L. Manual for the Vocational Preference Inventory. Palo Alto, Calif.: Consulting Psychologists Press, 1975. (a)
- Holland, J. L. The use and evaluation of interest inventories and simulations. In E. E. Diamond (Ed.), Issues of sex bias and sex fairness in career interest measurement (National Institute of Education Report). Washington, D.C.: U.S. Government Printing Office, 1975. (b)

- Holland, J. L., & Gottfredson, G. D. Sex differences, item revisions, validity, and Self-Directed Search. *Measurement and Evaluation in Guidance*, 1976, 8, 224-228.
- Horton, J., & Walsh, W. B. Concurrent validity of Holland's theory for college-degreed working women. *Journal of Vocational Behavior*, 1976, 9, 201-208.
- Johansson, C. B. *Manual for the Career Assessment Inventory*. Minneapolis: National Computer Systems, 1976.
- Johnson, R. W. Content analysis of Strong Vocational Interest Blank for Women. *Journal of Vocational Behavior*, 1974, 5, 125-131.
- Johnson, R. W. Relationships between female and male interest scales for the same occupations. Journal of Vocational Behavior, 1977, 11, 239-252
- Johnson, R. W. A "unisex" occupational scale for the Strong-Campbell Interest Inventory. *Applied Psychological Measurement*, 1978, 2, 527-532.
- Johnson, R. W., & Campbell, D. P. Basic interests of men in 62 occupations. *Journal of Vocational Behavior*, 1974, 5, 373-380.
- Johnson, R. W., & Johansson, C. B. Moderating effect of basic interests on predictive validity of SVIB occupational scales. *Proceedings of the 80th Annual Convention of the American Psychological Association*, 1972, 7, 589-590.
- Kincannon, J. C. Prediction of the standard MMPI scale scores from 71 items: The Mini-Mult. *Journal of Consulting and Clinical Psychology*, 1968, 32, 319-325.
- Kuder, G. F. Some principles of interest measurement. Educational and Psychological Measurement, 1970, 33, 205-226.
- Kuder, G. F. General manual, Kuder Occupational Interest Survey, Form DD (Rev. ed.). Chicago: Science Research Associates, 1975.

- Lamb, R. R. Concurrent validity of The American College Testing Interest Inventory for minority group members (Doctoral dissertation, University of Iowa, 1974). Dissertation Abstracts International, 1975, 35, 4161-A. (University Microfile No. 75-1216)
- Lunneborg, P. W. Interpreting other-sex scores on the Strong-Campbell Interest Inventory. *Journal* of Counseling Psychology, 1975, 22, 494-499.
- Lunneborg, P. W. Vocational Interest Inventory: Counselor's manual. Seattle: Career Decision Consultants of Seattle, 1977.
- Matthews, D. F., & Walsh, W. B. Concurrent validity of Holland's theory for non-college-degreed working women. *Journal of Vocational Behav*ior, 1978, 12, 371-379.
- O'Brien, W. F., & Walsh, W. B. Concurrent validity of Holland's theory for non-college degreed black working men. *Journal of Vocational Behavior*, 1976, 8, 239-246.
- Prediger, D. J. Do raw scores deserve a D minus? A reply to Holland. *Measurement and Evaluation in Guidance*, 1976, 9, 136-138.
- Prediger, D. J. Alternatives for validating interest inventories against group membership criteria. Applied Psychological Measurement, 1977, 1, 275-280.
- Prediger, D. J., & Cole, N. S. Sex-role socialization and employment realities: Implications for vocational interest measures. *Journal of Vocational Behavior*, 1975, 7, 239-251.
- Prediger, D. J., & Hanson, G. R. The distinction between sex restrictiveness and sex bias in interest inventories. *Measurement and Evaluation in Guidance*, 1974, 7, 96-104.
- Prediger, D. J., & Hanson, G. R. A theory of careers encounters sex: Reply to Holland (1976). *Journal of Vocational Behavior*, 1976, 8, 359-366.
- Prediger, D. J., & Hanson, G. R. Must interest inventories provide males and females with divergent vocational guidance? *Measurement and Evaluation in Guidance*, 1978, 11, 88-98.

- Prediger, D. J., & Noeth, R. J. Effectiveness of a brief counseling intervention in stimulating vocational exploration. *Journal of Vocational Behavior*, in press.
- Prediger, D. J., Roth, J. D., & Noeth, R. J. Career development of youth: A nationwide study. *Personnel and Guidance Journal*, 1974, 53, 97-104.
- Rayman, J. Sex and the single interest inventory: The empirical validation of sex-balanced interest inventory items. *Journal of Counseling Psychology*, 1976, 23, 239-246.
- Schlossberg, N. K., & Goodman, J. Imperative for change: Counselor use of the Strong Vocational Interest Blanks. *Impact*, 1972, 2(1), 25-29.
- Spokane, A. R., & Walsh, W. B. Occupational level and Holland's theory for employed men and women. *Journal of Vocational Behavior*, 1978, 12, 145-154.
- Strong, E. K. Vocational interests of men and women. Stanford, Calif.: Stanford University Press, 1943.
- Strong, E. K. Vocational interests 18 years after college. Minneapolis: University of Minnesota Press, 1955.
- Tilton, J. W. Measurement of overlapping. *Journal*. of Educational Psychology, 1937, 28, 656-662.
- Tittle, C. K., & Denker, E. R. Kuder Occupational Interest Survey profiles of reentry women. *Journal of Counseling Psychology*, 1977, 24, 293-300.
- Tittle, C. K., & Zytowski, D. G. (Eds.). Sex-fair interest measurement: Research and implications (National Institute of Education Report). Washington, D.C.: U.S. Government Printing Office, 1978.
- Zytowski, D. G., & Laing, J. Validity of other-gender-normed scales on the Kuder Occupational Interest Survey. *Journal of Counseling Psychology*, 1978, 25, 205-209.

	•		
<b>y</b>			
:			
<u>,</u>			

#### **APPENDIX**

This table is reproduced exactly from the supporting materials distributed as part of the presentation by Richard Johnson at the 1978 National Convention of the American Psychological Association.

AYALTHE	R-14(5MZ	I-MIZME	A-THEME	S-THEME	E-THEME	C-THEME	ACRICULTU	RE NATURE	ADVENTU
-тикчт	1.000	1.330							
-THEME	-057	•220	1.000						
-THEME	131	224	. 331	1.000					
-THEME	#361	*502	.050	101	1,000				
-THEME	.317 .416	.349	-,(31	.348	.658	1.000	1.000		
GPJCHLTURE Ature		*134 *361	- 426	•167 •306	002	7.047	.644	1.560	
DVENTURE	.540	, 376	.020	.045	.314	157	. 277	.084	1.000
ILITARY ACTIV.	,384	.235	128	. 201	.295	.374	.001	.027	. 30
ECHANICAL ACTIV.	. 404	.526	014	*051	*219	.408	, 258	-201	. 44
CIENCE	.522	.702	.043	•072	.074	.317	•117	.317	. 281
ATHEMATICS	<del>3,5</del>	.596	230	. JZ <del>V</del>	+225	-598	017	207	.21
EDICAL SCIENCE EDICAL SERVICE	.212	.474	138	.461	201	.301	.157	.369	.12
JSIC/DRAMATICS	015	.149	874	.342	,042	-1124	.111	.345	02
RI	1031	.163	.892	.305	032	129	.157	,497	07
RITING	012	.172	.810	.343	+104	059	•0+1	1292	00
EACHING	.035	.274	,381	+720	-134	140	.094	. 273	-,06
OCEAL SURVICE	021	.59a	.423	. 325	,258	.141	-164	, 259	-,039
THLETICS OMESTIC ARTS	•332 •021	.083	+.234 .375	•373 •527	.391 .198	+417 +202	.249 .174	,014	-,146
ELICIOUS AUTIFI	.100	,165	.254	413	.1+6	•241		. 2 6 1	- 0+1
UBLIC SPEAKING	.144	.247	.339	.477	.445	293	015	.033	. 244
AW/POLITICS	. 20♥	. 278	. 204	.399	. 442	.329	÷.077	054	. 282
ERCRACIO ING	,220	•C61	.133	.414	,872	• • 0 ?	4.038	.005	.103
LLES	.283	.027	-,127	.221	1854	.548	028	140	.239
ISINESS MOMT.		•148 •131	014	.389	,844	.722	035	0.0	-,069
ARMER	- 75 1	207	010	-, 420	.516 -,114	•783 •047	034	•01 <sup>7</sup>	.071
NSTRUM. ASSEMBL.	7 .241	082	720	-,249	132	. 392	009	235	.216
DC.ACRIC.TURK.	- ***5	.240	453	.213	.424	•521	+431	.212	1
IET IT LAN	174	.279	+177		,540	.589	.047	1255	4029
CLICE OFFICER	9 .767	-363	094	• 130	573	-524	353	152	. 671
RMT OFFICER	, ,383 2 ,448	.240 .274	-,511	.045	.413 .497	.474	073	122	.576
MYS.ED.TEACHER	·	.145	023 393	.355	222	.379 (293	-744	050	
CILLED CICITIS	F +287	107	541	-,363		1200			223
PRESTER	767	.375	151	4029	.150	+224	.444	.554	.340
W.TECH. (X~RAY)	£ .434	-460	064	.189	.219	.386	.164	. 294	.301
ERCH. MORTOFF.	0.6≜. ₪	.205	346	355	.009	:113	4322		• 376
AVY OFFICER	707	• 602	094	. 1 4 4	.502	.404	.057	003	.533
URSE, REGISTERED	3 .202	.481 .294	243	-+152	-,326		.208		
ARTOGRAPHER		447	167	+016	354	185	164	.120	.365
HERY OFFICER	± .755 □ .502	404	- 20*	.289	653	95	.024		4 9
IR FORCE OFF.	B +771		092	•133	, 178	1630	.105		-46
CCUP. THERAPIST	£ .447	. 4 2 0	.603	.569	. 252	212	.294	.501	. 2 2
NGINEER	f .460	+717	167	118	.237	. 487	. 047	0.0	. 34
NGINEER	E ,773	.730	078	004	, 238	445	.141 .070	+175 +138	.46
REMIST	£ .515	730	054	204	085	178		128	15
HYSICAL SCIENT.	= .00A £ .570	.758	-,249	085	624	301	<del>:</del> 16;	<del></del>	
EDICAL TECH.	£ 441	.753	201	079	.147	.384	.044	•170	. 21
HARMACIST ENTIST	f .555	.753	072	137	046	.215	170	1301	. 3 2
ENTIST	E .240	.404	1112	+064	040	.079	, 184	• 415	414
ENTAL HYGIENIST	f .153	.445	037	238	.05+	.132	.148	+342	<b>Q</b> 0 9
HYS, THERAPIST	E ,451	.784	.013	.407	279	. 4 4 2	. 273	<u> • 371</u>	.31
HYS I CAN	m .).9	•791	+24L	+127	-,044	0.9	201	148	.27
ATH-SCI.TEACHER	m .604 f .491	.835	001 548	059	.104	.378	.182	06*	.20
<u>ath-scliteacher</u> Iethilan	£ .491 £ .294	.611	.199	.422	.369	**0	960	. 357	.13
EDICAL TECH.	505	.825	005	.205	1 4 4	.427	.081	.310	.21
PTOYETR IST	E .45n	.778	047	.213	.314	.499	006	-181	. 23
DEPUTER PROCE.	£ .+3+	*102	181	104	251	,523			• 3 •
OMPUTER PROGR.	m .432	.765	044	-018	246	.519	007	•110 ••073	.03
ATHEMATICIAN	f .185	.231	-,21ª	423	420	162	-,135	- 663	-,27
KAIDITAMBRTA TEIDIEYH	=196 f .490	.231	211	410	-,674	074	.034	003	. 35
IOLOGIST	B - 140	.272	. 248	234	487	460	046 _	. 187	29
ETERIHAR LAH	7 .387	.446	091	519	-,268	207	.303	:3701-	
PTOMETRIST	.404	.475	442	384	.127	.322	054	065	. 2 2
RYSICIAN	186. 3	•710	-,087	247	-,183	e.036			-,29
CTAL SCIENT.	384	. 2 3 2	.437	167	-,591	* . 552	-,161	.009	.01
OLLEGE PROP.	f 134 m 154	.232	.754	•575	-,212 ,034	150	.038	-317	09
PEECH PATHOL.	1 .112	.367	024	-,484	-,459	330	04B	004	•07
PEECH PATHOL.	P - 272	.147	.373	245	-,731	594	007	. 2 9 5	-,29
TELDOLOGIST	f ,149	.499	.304	201	173	237	052	-104	•17
SYCHOLOGIST	B128	.311	.447	115	-,215		059	- 040	11
INCUACE INTERL	f 10A	•140	.347	419	424	481	102	.040	07
RCHITECT	B 287	, 7 <u>4 7</u>	.047	173	045	5 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	198.		
DVERY (SING EXEC.		242	.145	545	641	+.777	044	-,041	-,15
RTIST RTIST	f = 297	214	. 1 4 4	260	- 499	774	.024	.175	-,29
	f .136	.154	.782	.275	.009	221	, 184	. 430	.01
	n 274	112	.582	197	-,522	726	.041	.223	-,14
"STCIA"	9-149-	012	.772	199	- 222	315	.074	- 292	09
		043	. 68	034	-,244	367	054		- 4 4 3
USICIAN	a -+247 f -+252	144	77	.001	204	555	.047	.195	• 0 9

WARIABLE	R-THENE	I-MEME	A-THEME	S-THEME	E-THEME	C-THENE	ACRICULTU	RE NATURE	ADVENTURE
	1146	210	. 499	255	-,116	414	054	.075	1zŦ
INT. DECORATOR		102	.53B	+035	-1223	447	098	154	3.2
ADVERTISING EXEC.		*,157	704	.274	.135 085	245	.005 087	•150 •114	
The same of the sa	-, 422	1 95	.413	.374	010	₩.752 -4088	-,152	087	797
	375	.00z	.661	, 1 2 8	-,240	232	-,253	.144	-,469
REPORTER	343	-,273	.380 .490	→, 298 •092	-,350	594	102	067	075
KETUKIEK -	587 277	049	.717	,374	055	-,154	073	.138	144
ENGLISH TEACHER F	- · I B 4	044	.771	.478	.027	210	.074	• 274 • 732	-,584
MURSE, REGISTERED 1		.478	.099	.588 •529	.163	• 174 • 452	*207	.411	.346
LBIS'INCKALISI -	s +302	.007	.218	.529	094	.161	.035	.267	277
SCCIAL VORKER	-,026	•216	451	.605	173	179	103	1 20	- 103
***************************************	n = 126	.132 .068	.570 .667	.653 .614	.287	+129 081	-,104 -,024	•157 •240	-,103 -,133
Filipara	E -,107	.098	.560	.784	.147	860,	.077	.266	-, 146
YHCA STAFF	£ +043	.109	.431	• 750	457	201	.045	.132	•137 •351
Children	1027 101	.283	.484	.754 .788	-317 -400	182	110	.358	273
	000	,131	.344	.797	.266	*340	.101	. 2 8 8	227
SCH. SUPERINTEND.	123	+225	.200	.756	.591 .420	•524 •500	-,034	+.CS2	.113.
**************************************	.152 .081	•225 •151	,232 ,141	+511 •785	. * Z U	,497	014	055	.093
	£ .131 _	.155	. 908	. 831	.560	.367	.090	.180	.175
RECREATION LEND.	- 207	,165	.125 .167	,787	.311	• 497 • 195	+125 ++071	.113	C13
*******	f =.075 f =.055	.278 .128	.356	•725 •532	,333	-214	117	• 0 0 Z	.221
Market Street St	100.	7.129	.234	705	525	.298	.0.5	(00)	142
PERSONNEL DIR.	n +145	.155 .104	063	.523 .401	708 751	.472 .570	107 067	C.Q	.240
Market Market Livingson	m 4220			. 636	306		-103	*33A	
	£ +127	.104	.339	.584	• • 7 a	402	.047	+140	.296 .176
CH. OF COMM. EXEC. 1	n -057	~.011	.266	-502 -395	,678 1797	.439	- 157	124	235
THE PARTY OF	m .144	.047	.057	.545	726	.477	000	073	.302
LIFE INS. ACENT	f -144	.103	-127	.432	,752	+474	104	198	+174
Para r n u	£ -,022	.072 .025	.122 .498	177 .373	•094 •261	044	*.15l	.010	,514
	m174	.350	.102	.429	.767	572	043	064	• • 0 7
	n137	-,263	036	419	019	184	-,322	-,253	+211
I DVIC OCTOI	m .38# f .nzn	.488 159	~.150 394	•371 •074	.445 .587	4670 •522	.044 220	348	
BUTTE	m .0+4	106	+.036	+341	907	554	194	- + 2 3 7	•12e
CREDIT MANAGER	m +281	.213	057	.455	.843	798	129 028	110	.166
* Dividion Dividion Control	m .154	055 178	-,210	.191 .082	770	.605 .526		::	.239
MALE SUR	m -252	-,123	- 493	023	934	-530	.083	184	,003
PURCHASING AGENT	m .484	.184	- 267	+220	830	•735	142	078	.314
	.305	+6+. 101.	.306 405	+537 +124	613	.387 .742	- 301	- 422	. 206
ACCOUNTANT BANKER	m = 413	015	-,484	.001	,524	.762	215	377	-240
BANKER	m •331	-,174	436	•000	\$537	.580	-,251	431 405	.134
CREDIT MANAGER DEPT.STORE SALES	f +167	027 044	473 13*	.100 .385	,664 ,574	.740	035	035	240
BUSINESS ED. TCHR.	f 03#	-, 2 4 B	280	.332	.481	.638	163	370	135
	263	.125	034	.527	637	•75B	054	061 -131	-,rg5
ACCOUNTANT	£ .219	.188	-1241	-607	1599		226	367	• 257
SECRETARY	2 140	-492	102	.178	402	+416	- , 1 25	294	141
DENTAL ASSISTANT	£ -162	193	-,326	.315	269	+620	.144	.053	· C3 *
NURSE, LIC. PRAC. BEAUTICIAN	f +252 f -+202	.513 -,607	.061 561	-,173	023	017	039	-, 126	105
AOR	220	764	.570	.329	019	•117	.105	.491	275
IE PEROVERS	112	232 .003	-,488	692	-,538 -,011	-,302 +015	~.034 ~.034	146	208
TOTAL RESPONSES ) LP	+022 ,415	.471	.476	.545	.485	• 4 1 B	.177	.324	.362
OCCUPATIONS LP	.307	.219	,045	.253	.435	. 425	101	097	738
90 (	529	404	-,338 .432	-,554	-,476	626	200 .165	287	.141
SUBJECTS ) IP	.31*	404.	089	•091	-244	.251	.049	012	358
אַס (	934	-,597	314	464	419	540	-,219	-,395	199
ACTIVITIES IF	- 292	.336	C85	+579	, 404	+320	+036	00	043
ACTIVITIES   IF	.199 475	400	269	578	-,594	544	148	315	-,293
) LP	.25*	. 354	511	- 474	326	2 10	.114	.325 .018	·233
AMUSEMENTS   IP	.184	.095 394	058	*096 *****	-,485	•252 ••443	-,170	- 295	-,248
TYPES ) LP	105	257	, 408	417	280	-167	.048	.228	.198
OF IP	.034	111	217	097	034	1217	- 002	099	C73
PEOPLE ) DP 1 LP	140	[83 -093_	-,22* ,154	403 013	302	234	088	155	-142
PREFERENCES -P	-,004	044	040	~.045	004	.027	097	117	C.1
RP	040	017	072	.120	.384	-087	051	.104	072 .171
CHAR- YP	.151	*197	-,129	084	.002	-081 -028	.032	033	004
ISTICS NP	-,230	176	123	212	202	115	-,078	095	200
	- '	•							

WARIABLE HILITARY ACTIV.	HILITARY VITOL	ACTIV.	SCIENCE	MATHEMATICS	SCIENCE	SERVICE	HRISTE DRAMATICS	ÁRT	WRITIN
RECRASTICAL AUTIV.	1.000	1.000							
SCIENCE	.252	.577	1.000		•				
HATHEMATICS	.278	474	.403	1-000					
MEDICAL SCIENCE	+217	.283	. 6 4 0	.325	1,000				
MEDICAL SERVICE MUSIC/DRAMATICS	.224	.150 083	.993 002	.211 227	•721 •127	1 c 0 0 0 + 1 7 9	1.000		
ART	C3* -,144	-,035		-,23	- 117	1149	.757	1.000	
WRITING	- 124	007	001	213	104	.059	.476	, 63 T	1.00
TEACHING	050	013	.cat	025	.175	+ 250	.350	,351	, 9 2
SOCIAL SERVICE -	.020,	129	069	174	,187	1358	. 510	404	- 44
ATHLETICS	.324	.249	.143	. 2 4 6	*53B	• 234	-,247	-,244	24
DOMESTIC ARTS	.042	057	370	039	173	·41B	.348		.20
PUBLIC SPEAKING	.222	021	.107	085	183	040	,304	175	. 97
LAW/POLITICS	.149	•116 •156	.130	.100	204	1040	.141	.054	38
TERCILORD TAC	215	.148	050	1172	131	+171	.128	.141	+17
SALES	.250	217	045	.200	061	.050	113	124	04
BUSINESS HOME.	.318	. 284	.052	.243	.158	+149	= .022	01*_	.07
OFFICE PRACTICES .	.263	.126	*107	.313	.177	+373	.024	-014	• 31
Parmer Instrum.assembl.	4131	.277	.017	.213	154	103	-,455	*••0 <del>*</del>	80
100 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	<u> 2 6 6 </u>	.319	-107	-342	047	-045	431	583	-,44
DIETITIAN	D .354	•590 •077	+314	.378 .204	,265 1457	+262 +427	.234	-227	.13
POLICE OFFICER			.337	.354	354	294	127	141	-,09
HWY . PAINOL OFF.	型 <u>.400</u> 四 <u>.557</u>	.652	+321	.491	.235	1182	-,975	-,479	•.50
ARMY OFFICER	f .248	.344	1218	.22	1148	117	-,119	182	+17
PUYS.ED.TEACHER	1 340	. 2 Da	.173	2 6 1	321	404	312	351	-,47
SKILLED CKAFIS	.221	. 454	,310	.36+	.025	.010	-,524	-,446	-,67
FORESTER	m .255	.445	.440	.317	.3:4	. 205	<.203	095	23
RAD.TECH.(X-RAY)	2 .310	.926	.726	.475	<u> </u>		046	D46	- 14
MANY OFFICER	B .217	702	,335	.312	.041	021	-,386 -,127	267	-,42
	- 4334	•729 •059	•581 •427	•679 •131	•372 •751	.204	.282	251	.15
		•131	,382	.135		442	085	015	30
- 4	D .344		.736		411	211	200	158	21
ARYY OFFICER	<u> </u>	140	.350	. 5 4 5	300	.172	-,226	264	10
AIR FORCE OFF.	502	.313	.495	.707	•413	.276	-,:+3	104	10
	f .126	.377	, 456	+173	+412	.544	502	. 4 4 7	• 43
	£ .312	.755	.778	•772	.382	*4**	215	193	17
LIU LIEUX	177;	. 8 6 1	.793	. 498	+417	. 258	326	099	13
	1 .154	-428	,810	.5 7 2	397	100	.037	.038	-00
	E -, 238	.148	.512 .840	.488	123		-,263	2 2 3	-,52
MEDICAL TECH. FBARMACIST	f ,295	.471	110,	. 4 4 4	700	499	205	171	-,23
DENTIST	£ 207	. 614	.019	499	450	. 390	1 4 1	107	14
DENTIST	B .098	. 254	,533	243	.710	.527	•102	*12Y	50
	f ,17*	.097	.452	.174	,755	.774	.000	.050	4
PHYS, THERAPIST	f ,336	.415	.765	.544	, 5 4 7		*038	.025	05
PHYSICAN	m -114	381	.785	1778	714	-529	-,211 -,034	- 035	14
	t .264	.454	.894	.69B	,574	.432	-,531	025	53
MATH-SCI.TEACHER	. ,,,,,	.591	.614 .558		<u>•364</u> •577	.550	-202	,275	
DIETITIAN MEDICAL TECH	_ •••	.240	.571	.413	239	1474	011	015	<b>*•07</b>
MEDICAL TECH. DPTCMFTRIST	301 282	.502	.757	.657	733	.584	070	077	78
COMPUTER PROGR.	2 ,303	.724	757	.771	390	.190	240	181	17
	307	7 4 1	. 8 2 1	.783	435	244	071	077	07
34117-3411-316	£076	.320	.474	.371	.026	25.5	*,247	243	18
CANTITUDES TO CALL	398	075	.299	.108	004	-,111	.035	.101	-05
	1 .120	+417	.497	•540	.229	080	281	280	20
BLOCOCISI	225	112	+327	058	179	•127	. 227	4282	12
A C I C X FUVY PYC.	031	.399	.579	• 252 • 575	+372 +414	.087 +143	153 444	117 957	17
or togging t	230	-507	.422	.470	165.	,200	-,13*	152	10
PHYSICIAN SOCIAL SCIENT.	1 - 448	7.312	004		-,054	130	.382	.362	- 4 6
COLLEGE PROF.	1204	105	106	-,188	1111	144	.337	.247	. 97
COLLEGE PROP.	119		- 0+0	344	. 289	229	.736		.71
SPEECH PATHOL.	-,126	.218	.441	+182	.053	257	044	083	. 05
PEECH PATHOL.	387	187	. 184	185	029	057	1344	.345	• 3 0
SYCHOLOGIST	£175	.20*	447	-122	230	117	.158	-174	35
	401	115	.184	21*	.191	.057	.554 .291	.582 .254	
ANGUAGE INTER	2331	054	+111	-,170	-,057	-,30f	440	,454	. 3
RCHITECT DVERTISING EXEC.	P151	-,134	-,311	249	-,371	641	010	043	
ARTIST	1 400	243	224	403	-,327	481	.072	.127	.00
	,529	-,394	233	521	-,212	204	.301	.443	32
	(133	.078	004	257	.022	.004		,626	• 5
PHOTOGRAPHER	1 4	-,298	-,123	441	-4154	200	.504	7549	• 97
MISTOTAN	2 - 173	217	102	348	-,010	,044	.05	.658	
NUSICIAN	-,574	-,214	-,123	340	- 073	028	. 682	. 4 7 4	. 4 !
INTERTA LNER	-,287	308	241		100	092	.474	1414	. 5
	<u> </u>	-131	ــلېږـــ	- 257	3-0	- 1439	. 367	.549	
	-,373	400	518	58a	-,279	+ · 1 2 8	.528 .411	.623	7
DVERTISING EXEC.	205	- '210	-,351	535	011 173	170	584	.534	7,
ANGUAGE TEACHER	£ -,737 £ -,138	-,520	- <u>,384</u> ,084	-,438	-,054	-,094 -,243	.418	.383	<del></del> ;
LIBRARIAN	=289	332	078	332	053	.001	403	593	
IBRARIAN EPORTER	1 - 359	-,367	- 364	-,575	-1331	- 192	, 274	.261	<u>.5</u> c
	334	-, 6 10	- 42	-,453	-,200	4.197	.432	.385	
LZPORTÉR							.424	.549	,

STRIBAL		RILITARY	HECHUTICAL			MEDICAL	MEDICAL	MISTO		
		ACTIV.	ACTIV.	SCIENCE	MATHEMATICS			DRAMATICS	ART	URITING
NGLISH TEACHER	В	15t	273	~.224	509	,012	- 0 6 4	.285	.268	200
MRSE, RECESTERED		.197	.015	• 375	• 0 7 8	•773 •747	.439 .481	.077	.077	.240
MYS.THERAPIST	B	.388	.519 335	058	.369 157	365		.305	.284	.174
MRSE LIC PRACE	- <del>-</del> -	-040	093	,000	135	240	.226	.412	.345	,547
OCIAL WORKER OCIAL WORKER	<u>,</u>	04:7	- 194	- 077	-,229	176	.215	.520	. 468	+471
RIEST	-	-,095	- , 294	122	331	.110	-199	.670	.569	.489
IR CHRIST . ED.	í	004	206	087	-+191	.121	+27+	\$577	484	•577
WCA STAFF	f	.044	040	102	141	.154	+172	.404	. 345	495
CINISTER	70	.045	065	.063	095	+300	.346	.433	.528	45 9
ELEM, TEACHER	<b>a</b>	.121	.075	.168	+064	405	. 434	,449	.378	,458 ,334
LEM. TEACHER	£	.144	078	.032	•030	.208	.454	,212	.132	.298
CH.SUPERINTEND.	_ =_	• 50 •	.046	+026	•102 •112	- 214 192	- 255		-115	411
PUBLIC ADMINISTR		-194	DOS	054	.073	274	.308	.194	.137	.241
WIDANCE COUNSEL	. 🕦	.170	-015	041	016	- 7 2 1	.311	.404	.340	447
ECFEATION LEAD.		297	+071	.603	•077	254	-307	.215	-138	293
ECREATION LEAD. WIDANCE COUNSEL		.000	154	.006	071	210	. 232	.429	.382	.541
OC.SCI.TEACHER	Ť	.008	**140	078	077	.101	.036	.290	, 229	+551
OC.SCI.TEACHER	- E	.114	145	305	253 .	• G 4 5	.121	.220	.170	.342
ERSONNEL DIR.	-	.182	.079	033	+041	.172	1117	.187	.127	+708
EPT. STORE MOR.	. 13	. 299	.153	012	•174	187	•161	051	082	014
OME ECON. TCHR.	f	.075	110	061	034	.166	437	.377	470	.222
LIGHT ATTENDANT	£	.172	•00Z	-,074	<b>→•</b> 000	,225	.350	.370	.347	
H. OF COMM. EXEC.		-197	006	117		,084	-058	,238 ,C58	.154	• 4 4 <u>3</u>
ALES HANAGER	-	• 1 B &	.073	167	-052	174	•027	.074	010	.151
ITE DIS. ACENT	g. Í	.327	.034	-,074	+100 +087	107	•174 •016	.117	.027	,257
IFE INS. ACENT		.189	001	069	046	104	415	.025	070	371
ahyer Ahyer	-	077	-, 286	172	228	042	030	435	.337	.495
OMPUTER SALES_	_	,287	.300	.211	.345	. 276	.172	.074	.009	-179
NVESTM. FUND MCR		-,223	297	325	176	-,312	495	057		.079
HARMACIST	. 19	,343	.352	. 441	. 444	.428	.405	110	-,163	~.105
UYER	ŕ	· 20A	.037	~.154	.175	044	•021	-,301	373	-,294
LYER		.197	.017	287	.043	-,054	+G18	-,008	022	.033
REDIT MANAGER	<b>B</b>	.348	.245	,114	.347	,230	.234	+.037	-,985	-060 192
UNERAL DIRECTOR	12	.273	•097	130	144	162	. 266	-,120	147 437	-,172
EALTOR	-	.255	.163	190	•200	- 677	057	-,577	574	- 1665
LRIBUSINESS MGR		. 349	.228	002 -159	.332 .387	.000 .20:	.104	240	-,242	217
TREHASING AGENT	- 0	.357	.454	-134	•149	581	.489	272	.238	.3.7
HIROPPACTOR CCOUNTANT	_	.310	.264	.139	574	051	.00*	-,352	440	₩.235
ANKER	<b>13</b>	.293	.230	.020	. 464	091	047	426	481	310
ANKER		.222	.048	161	+243	-,072	111	-,463	462	261
REDIT MANAGER	f	.337	.203	·•012	.380	001	.039	385	478	324
EPT. STORE SALES	. Ē	, 329	-048	059	.226	800,	• 369	044	043	-,109
USINESS ED. TOHR	. £	.192	++051	297	+142	-,221	016	-1103	-,2a0	[0]
USINESS ED.TCHR	· u	.297	-208	.025	+270	178	1239	017	032	.047
XEC. BOUSEXEEPER	_ 4	.317	•171	.111	-300	-,058	. 475		399	212
CCOUNTAIN	£	,262	.415 179	.245 544	-645	-,135	165 025	028	023	041
ECRETARY ENTAL ASSISTANT	f	.095 .347	-144	.237	-375	434	•670	203	-,227	383
URSE, LIC. PRAC.	<del>-</del> -	4337	.205	504	- 343	675	.830	135	085	017
EAUTICIAN	£	.047	177	501	078	- 377	101	4.411	373	458
AOR		.047	.177	.643	.344	572	.415	. 488	.461	.555
12		135	010	023	.028	-,261	265	445	-,348	575
OTAL RESPONSES		+.025	024	002	110.	.004	004	024	019	003
> L?		- 224	. 105	,313	.167	436	<u> </u>	, 412	,432	452
OCCUPATIONS   IP		.220	4 2 9 1	,201	.246	, 2 ; 5	.255	.051	.054	.052
ים (		- + 371	433	369	+.315	- 457	-4472	303	318	-,327
SCROOL J.P.		.157	.258	. 461	. 375	456	- 345	. 363	- 364	- 415
SUAUECTS } IP		.176	-150	-040	-145	.063	.109	075	075 245	110
) DP		310	+.377	477	-+494	-4474	-,422	. 356	.346	. 179
ACTIVITIES   IP		184	• 218 • 227	,195	+177	331	-059	092	C78	
ACTIVITIES } IP		-,243	434	301	271	372	421	-,234	241	-,254
) LP		147	. 209	.214	.075	4301	.311	. 495	. 444	• 415
ANUSEDENTS   IP		179	•197	.110	.163	.044	.071	061	070	-1021
DP		277	344	294	236	-,329	333	364	215	304
TTPES ) LP		106	.042	.142	012	.231	. 280	. 4 2 0	.364	. 341
OF   IP		-064	.029	049	.D47	048	624	-,201	-,233	
PEOPLE DP		-,220	093	11*	050	198	-,322	267	199	148
) <u>L</u> P		-,093	.015	.053	030	.048	.026	-130	-154	.091
PREFERENCES } -P		05	*077	013	•049	-,088	045	032	070	C45
		.013	<b>+.</b> 047	030	040	074	.081	091	054	
יטא ל										
CRUR- YP		042	.182	122	1005	123	079	182	£40.5	.252
יטא ל			•182 •052 ••271	016 121	.005 .054 058	-,123 -,636 -,103	033 050	-,12)	093	157 127

ATALABIC	:EACHING	SERVICE	ATKLETICS	ARTS	ACTIVITIES	PUNLIC SPEAKING	LAW POLITICS	MERCHAN- DISING	SALES	BUSINESS	OFFICE PRACTICE:
TEACH ING	1,000				A					HANACEMENT	
ATHLETICS	.574	1.000	1.000								
DOMESTIC ARTS	•110 •391	.521	.054	1.000							
EXCIDIOUS ACTIV.	.32*	. 459	+174	1376	1,000						
PUBLIC SPEAKING LAW/POLITICS	.324 .252	•413 •332	•177 •144	•071 -•027	.2.1 .083	1.000	1.000				
MERCHANIO TAIC			. 296	-312	184	. 114	- 38 q	1.000			
SALZS	047	+102	.334	.049	.074	.304	.304	.742	1.000		
BUSINESS MONT.	.185	.240	,373	- 152	.147	. 467	.121	.811 .549	,737 .450	1.007	1,000
FARMER .	,220 =485	512	.258	-402	.294 183	553	450	187	.054	-,112	014
INSTRUM.ASSEMBL.	£3e#	378	.329	100	100	424	332	4073	,264	+151	.375
DISTITIAN .	-C34	**012	,550	. 643	.125	-018	4208	,282 ,595	• 373	.552	.443
	.347 .058	.474 .124	.274	.052	.429 .175	.286 .321	.161	1494	.473	,534	. 293
	173	182	.623	176	,005	.021	4   1 4	. 242	• 420	• 2 9 6	.222
ARM OFFICER PHYS.ED.TEACHER	.05  .120	022	.257	-,384	-,117	.599	734 052	.351 .147	+431 +184	.563	4088
	120	-139	.748	-105	-,143	047		104	109	011	,035
PORESTER :	043	144	.337	.037	.044	086	082	.011	.080	102	.044
EAD.TECH.(X-RAY)	031	.059	.346	-142	.174	•001	.027	-100	- 092		047
	324	438 070	.124	243 047	-,215	-,144	-,200	101 .357	.389	-,514 ,526	.288
	2291	.382	147	.444	.372	.047	011	.050	091	.020	. 284
PARTOCRAPHER	132	- , 2 2 6	.045	•004	,037	497	-,53%	365	- 3.2	407	107
CARTOCRAPHER C		154	.315	045	.057 .138	. 934	.159 .493	.204 .514	.524	.729	+292 +397
AIR FORCE OFF.		079	,557	-033	112	+253	258	~~ <del>;</del> 343	348	<del>: \\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\</del>	.303
CCUP.THERAPIST	.519	.507	.127	+470	,331	.291	.200	.250	•020	4176	+175
ENG LIVE EN	0**	315	,211	225	÷,059	•074	-173	-075	-172	-220	-140
ENGINEZR CHEMIST	025	178 326	.228 035	047	-,039 -,103	•07 <b>•</b>	*108	-,282	138	- 4 C 8 4	171
	049	297	450	176	127	285	-,227	-,471	4 20	&D1	-,34;
MEDICAL TECH.	134	+,248	.223	107	•005	129	019	058	.042	.074	.130
PHARMACIST	-,180	224	•211 •131	049	-020 -110	047	.025 .047	•041 ••104	.094 023	002	024
DENTIST S		007	.073	-0130	.010	040	045	105	125	125	028
PENTAL STOTEMENT	.071	.155	.253	.318	133	077	094	•05*	4.040	000	•170
PHYS.THERAPIS:		. 205	• 471	.232 .098	,310 ,124	-153	.044	-145	114	117	042
PHYSICAN MATH-SÜLTEACHER "	• • • •	025	.031 .225	•050	1134	•001	053	030	021	.05a	.147
ATH-SCI.TEACHER 1	-,097	28 -	.369	153	.003	137	024	-041	-205	• 2 3 5	, 359
DIETITIAN	274	284	.224	,583	249	.219	.107	.350	-030	.329 .136	.395
REDICAL TECH.		.052	.212	•101	+235 +145	+05 <b>†</b> +135	.103	1207	-174	.742	.257
COMPUTER PROCE.	CB3	271	. 2 2 8	157	035	+043	.171	• 1 3 8	-१४५	.247	.210
COMPUTER PROGR.		150	.172 243	041	.059 -,287	-125	.191 082	+114 ++577	.14S	369	3+5
MATHEMATICIAN F		295	5+1	185	- 222						432
PHYSICIST		-,507	004	-,483	765	091	.046	÷,353	154	-,143	263
BIOLOGIST		-,137	-,532	•053	031	357	373		-,732	677	353
VETERINARIAN	-,376	-,558	074	-,368 -,374	318	125	187	-,138	248	309 - t 25	397 <b>PC</b> 0.
OPTOMETRIST F PHYSICIAN		345	001	-,333	- 147	C37	.044	384	-+231	- 215	321
SOCIAL SCIENT.		.049	431	035	-,149	-+032	023	509	644	*.55B	
COLLECE PROP.	· - ·	-107	369	-+240	+,124 .307	.383 .458	.349 816.	239 .107	303	-,198 -326	5)6
SPEECH PATHUL.	-1164	-,464	178	518	-,294	078	- <del>-:313</del>	588	-, 439	-,433	
SPEECH PATHOL.		078	581	059	113	194	203	700	772	721	·.514
PSYCHOLOGIST 1	.579	150	262	304	-,348	+161	,282	255	278	-,104	45
PSYCHOLOGIST PLANCUAGE INTER.	.313 072	.276 -,233	443	309	347	+229 +014	.207 .078	177 425	434	273 410	m.520
	006	-,079	361	4077	-,043	- 015	049	135	269	- 257	25
DVERTISING EXEC.	197	356	-,267.	486	-,400	-137	4214	044	303	042	494
	251	-,35a 021	-,529 -,535	344	378 158	251 253	-,227 -,318	*,425 *,599	551	652 738	-,799 -,573
		.311	-,216	.333	118	•203	.090	1117	147	~~~ <del>~</del>	- 167
LRT TEACHER E		.037	578	.030	144	042	130	435	541	590	576
ersterva [	. 276	159	-,341	.303	216 017	-162	002	144	324	762	187
KUSICIAN ENTERTAINER		177	-,515 -,320	166	073	113	-,025	110	-,274	. = . 323 = . 287	37
NT. DECORATOR		080	502	024	-,174	028	017	009	125	180	390
NT.DECORATOR		.226	-,543	+322	.041	079	20		**240 •035	- 4267	-,15
DVERTISING EXEC.	• 271 • 517	.391	-,187 -,330	+ 225 - 388	,081 173	,445 ,287	,313	.251 .041	147	00	139
<u>Anghage Teacher</u> Ibrarian		.097	471	058	003	-376	241	=.005		•010	126
IBRARIAN T	333	4273	410	.241	-144	144	.001	071	320	203	4.02
EPORTER E	-,020	048	157	-,239	-,249	-173	.144	243	326	354	-,366
EPORTER ** NCLISH TEACHER **	.182 .499	.238 .474	-,372 -,287	+057 +244	-,034 .182	+195 +478	.148	.170	072	-0357	017
NGLISH TEACHER A		540	170	.351	299	432	.295	•122	111	• D 1 7	016
TRSE, REGISTERED 6	.374.	.514	*574	. 4 4 3	. 404	+172	+124	147	.014	.118	. 26
MYS.THEPAPIST		.328	.552 .047	270	,349	+233	123	. 236 . 18 T	-173 010	.292 .075	,317
DOSFILC POACT S		.538	020	1240	,244	010 24	.646	+339	-149	- 1075 - 370	.139
OCIAL WORNER P	· · · · · ·	714	076	.374	*310	. 4 2 2	,553	. 345	.078	.334	. 169
PRIEST C	520	781	144	.407		. 495			072	.047_	,054
	+727	.741	.004	.479		. 784	·+ 3 Q 5	.221	023	.176	.217
DIR, CHRIST. ED.		.727	.214	338	.310	.700	,544	. 424	.285	494	.201

CORRELATION MATRIX FOR TOTAL SAMPLE

	. <del></del>	SHEEKU.		DUMEST IC	id:Luctors	PITHLIC	LAH	MERCHAN-		BUSTRESS	OFFICE
ATELVOPE	TEACHING	SERVICE	ATHLETICS	ARTS	ACTIVITIES	SPEAKING	POLITICS	. DISINĈ	SALES	MARACEMENT	PRACTICES
MINISTER 3	.415	.796	.03*	. 454	,539	. 434	• 573	. 344	.104	244	.73F
ELEM. TEACHER	79	•742	• 2 2 9	,507	, 4C6	.453 .225	.371	.346	106	.374	.34.
ELEM.TEACHER É	.443	.428	.144	.440 .312	• • 1 4	.645	584	.575	.407	+289 	.534 .470
SCH_SUPERINTEND_ P	.553	*414	.190	145	163	780	767	.501	449	,720	, 3 + 8
PUBLIC ADMINISTR.	.532	480	.481	391	.413	.500	.441	.\$75	.362	50	463
RECYFATION LEAD. F	.578	.721	.381	. 447	• 9 2 3	.457	.503	,593	. 345	595	195
BECREATION LEAD. B		-,544	,582	*327	.412		*231,	- 2845	- मह	. 676	• > 1.7
CUIDANCE COUNSEL. 1	-447	.728	.047	.340	.317	+425	.554	.347	1111	.341	.215
SOC. SCI. TEACHER . E	.574	.519	4034	.144	. 191	-714	.770 .541	.349	•160 •405	.374	- 114
SOC. SCI. TEACHER		, 646	.276	,289 ,187	.307	7592 7686	.641	.495	.524	759	
PERSONNEL DIR.	.304 .097	.428 .235	,5)9	.123	135	. 444	441	.734	. 631	747	. 195
DEPT. STORE MOR. B	170	581	135	.036		-103	632			. 25 D	• 773
BUNZ ECON.TCHR. F		.515	.345	.505	. 287	.372	.319	+714	457	.541	6 5
CH. OF CONST. EXEC.	.307	. 424	.193	•1 <sup>0</sup> 2	,171	.749	.721	. 676	528	.727	.335
SALES MANAGER M	+122	.288	.347	+121	.113	.587	525	.749 .474	4623	•7•2	
LIFE INS. AGENT		. 404	.521	.205	.278	• • 0 6	.534 .711	.473	.642	•704 •747	.368 .292
LIFE INS. AGENT T		117	320	766	-,132	1544	-:45:			1137	
LAVYER T	-, DS 8	.464	111	154	140	498	470	.294	.123	.270	. 25 :
COMPUTER SALES		270	.373	101	163		. 4 2 4	. 457	.631	.734	. 3 5 4
INVESTM. FUND MGR.		-,326	-,268	-,317	454	+129	.217	.017	. G 2 A	.C25	1 . 1
PRARMACIST =	.094	.174	- 721	.223	.237	+147	149	. 401	,534	.578	.551
BUYER		015	.304	041	.015	-149	195	1584	1018		
BUYER		.243	.325	.17E	,114	.375	.370	810	.726	.402	4-74
CREDIT HANACER		• 2 6 C	.383	•1*2 •2*2	228	.307	244	.744	,710	.731	.54:
FUNERAL DIRECTOR S	217	094	.446	098	-, D22	.223	.241	.440	+7+7	.478	+342
AGRIBUSINESS MGR.		-,214	444	034	040	144	125	.347	.544	.439	. 472
PURCHASTING AGENT		.024	504	•03t	104	271	.341	.731	.757	.004	, 4 7 7
CHIROPRACTOR B	.273	454	.276	1269	, 262	.527	490	.462 .539	.358	,5C4	4277
ACCOUNTANT =		041	.909	-,113	000	4297 4119	.407	479	571	• 4 6 D	. 655
BANKIR	147	172	.388	149	044	+241	.333	.505	,547	.617	.373
BANKER E CREDIT HANAGER	-,145	044	357	055	021	.215	.280	. 40*	.495	.753	* 4 D R
DEPT STORE SALES		. 236	.315	.451	418	.103	.051	.629	.55*	.580	,827
BUSINESS ED. TCHR.	•244	.200	.262	184	155	• 172	,195	.536	.735	.584	4717 +693
BUSINESS ED. TCHR.		.345	.392	.202 .549	.236 .522	.289	208	.621	967	.863	.7 <u>4</u> 1
EXEC. HOUSEKEEPER	272	-,150	.194	284	-,037	4182	290	.307	1428	. 493	155
	221	.168	111	300	125	.00	004	593	. 457	.464	. 447
DENTAL ASSISTANT	.003	.144	4919	.347	. 285	054	081	• 363	.327_	.371	-711
RURSE LIC . PRAC.	247	.382	.311	.469	,562	-114		51.9	-124	.237	.513
	376	193	. 253	.047	-,054	409	-, 547	.098 057	.197 198	.026	.184 .072
AOR	• 447	423	154	-250 -0325	-,308	-,788	-::::	-,518			
12	-,484	.0)8	201	014	-,024	.004	032	018	024	003	.005
TOTAL RESPONSES	.423	454	.254	,300	.307	+472	.429	. 497	. 245	. 4 4 8	.372
OCCUPATIONS IP	.150	.150	.234	-170	, 175	+125	.177	280	-398	.404	. 4 2 4
) DP	-,372	418	361	-,384	3+0	455	417	437	508	430	594
SCROOL ) LP	.320	.347	.123	-348 -044	.315	•030	.352	+184	+012 +263	.176	212
SUBJECTS   IP	.014	-,342	.175	-,342	,100 -,384	-,274	-,345	371	-,254	379	42;
אַק ל אַר כ	-,310 PBC.	504	273 .172	.473	+328	482	. 385	.384	.104	.373	. 2 - 1
ACTIVITIES 12	027	037	-147	011	.058	+030	,050	.180	.243	• 2 1 1	•212
DP	330	430	310	424	747	474	404	541	-,420	541	4 • 4 4 0
) LP	.304	.393	.198	1125	. 157	4109	.306	.330	1119	.271	4164
AMUSEMENTS IP	*D29	.029	157	.D03	-,127 -,426	-088	35#	447	320	.223 -,442	.204
) DP	288	364	318	-,346	.281	435 -327	.256	.224	.052	.177	-115
TYPES LP	113	097	.025	D85	034	140	107	035	.076	352	:033
PEOPLE DP	232	345	201	251	-,324	221	175	228	159	240	185
) LP	.091	032	034	-026	05*	019	041	- 055	- 172	122	137
PREFERENCES -P	-,096	D47	-001	033	.023	034	-,042	017	.070	003	.03Z
J RP	.059	•193 •199	.011	•027 •091	.029	•075 •350	.115 .240	.124	.074	.130	017
CHAR- YP	094	086	049	047	-,025	132	010	034	C 6 2	-034	1017
ISTICS NP	144	134	130	052	074	257	170	130	155	170	.010

Note. The data are based on 1,134 male and 1,044 female first-year students at the University of Wisconsin — Madison.

			•	
		•		
<u> </u>				
i				
I ;				



